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Estimates of queer kinship
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Queer family matters: Estimates of queer kinship in the United States

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Abstract

Queer family demography largely centers the nuclear family, with limited attention to extended kin. We provide a more comprehensive picture of queer kinship in the United States, specifically aunts/uncles, cousins, siblings, nieces/nephews, children, and grandchildren. Combining demographic kinship modeling techniques with survey data that directly identifies sexual minorities, we estimate both the individual-level number of queer kin and the population-level share with at least one queer kin, focusing on age, sex, and cohort dynamics. Findings indicate the number and composition of queer kin are shaped jointly by queer identification patterns, population and kinship network age and sex structures, and past demographic rates that determined kinship network size. We find substantial cohort differences: younger cohorts have more queer horizontal kin whereas older cohorts have more queer descendants. A majority of U.S. residents has one or more queer kin, driven by (bisexual) extended kin. Beyond providing estimates in a context where no comparable data exist, this study advances a queer kinship demography that situates queer people as members of extended family networks. This study offers a mechanism linking exposure to queer people to demographic trends, with implications for broader acceptance of sexual minorities.

Keywords queer | kinship | family demography | kinship matrix model | sexual identity

*“...sexuality is embedded in kinship in ways that everyone knows but many hesitate to speak.
Where else would babies and fathers-in-law come from?”*

– Kath Weston (1997) *Families we choose: lesbians, gays, kinship.*

Introduction

All humans are embedded within broader kinship networks, though the structure of these networks varies over time and context (Alburez-Gutierrez et al., 2022). A primary focus of the kinship demography literature has been to identify and understand how variation in the structure of kinship networks has implications for individuals across the life course (Dixon, 2024; Hünteler, 2022; Ruggles, 1987; Verdery & Margolis, 2017). More recent work has focused on heterogeneity in the composition of kinship networks, such as the implications of educational attainment of family members for caregiving (Adhikari & Alburez-Gutierrez, 2025; Jiang, Liang, et al., 2025). Far less work has considered how queer sexual identity—here defined as those who identify as lesbian, gay, bisexual, or another non-heterosexual identity—is distributed within kinship networks, or by extension what share of the population has queer kin.

Even within scholarship that examines the family relationships of queer individuals, queer family demography in the United States largely centers the nuclear family. This literature highlights the unique experiences of same-sex couples and parents, or positions queer children in relation to their parents and siblings (Bosley-Smith & Reczek, 2024; Furstenberg et al., 2020). Such emphasis on family formation, childbearing, and the parent-child dyad subjects queer people to norms that privilege heteronormative nuclear families and limits understanding of broader kinship dynamics that include queer people (Allen & Mendez, 2018; Weston, 1997). Moreover, due to data limitations, many of these studies indirectly measure queer sexual identity by using the sex composition of couples or the identity of only one partner. These studies thus provide conservative estimates of the population identifying as queer as they exclude single queer people or those in different-sex partnerships (Julian, Manning, & Kamp Dush, 2024; Julian, Tessler, et al., 2025). In sum, the literature is presently limited to same-sex partnership and parenthood, leaving a gap regarding the demography of queer kinship.

The omission of queer kinship from the demographic literature is concerning because queer people represent a sizable population in the United States (Gallup, 2025; Silva, 2025). To address this gap, we advance a queer kinship demography that focuses broadly on individual and population exposure to queer kin, including aunts/uncles, cousins, siblings, niblings (nieces and

nephews), children, and grandchildren. Using a mathematical demographic model that produces ego-centric estimates of living kin including kin type, age, and gender for an average person we name “Focal” and survey data on queer identification (Caswell, 2022; Caswell & Song, 2021), we generate counts of queer individuals in kinship networks representative of average U.S. residents under rates of queer identification pooled for the 2021-2024 period. These estimates are then discussed with respect to the age, sex, and relationship of kin. Finally, we produce novel demographic estimates of queer kinship in the population, that is, the share of people with one or more queer kin.

Our findings indicate that there are substantial cohort differences in the number and composition of queer kin that reflect the interaction of several social and demographic processes: age- and sex-specific patterns of queer identification, kinship network age and sex structure, and the historical fertility and mortality regimes that shaped present-day kinship network size and population age-sex structure. Whereas younger cohorts have more queer aunts and uncles, cousins, and siblings, older cohorts have more queer nieces, children, and grandchildren. These counts are patterned by the age and sex of kin, with higher counts of queer female as opposed to queer male kin, particularly among younger cohorts. We also find that a sizable majority of the population has at least one queer family member, driven chiefly by queer extended kin.

Population trends in queer sexual identification constitute one of the fastest growing compositional changes in the United States and the implications for kinship network composition are evident, analogous to the impacts of educational expansion (Adhikari & Alburez-Gutierrez, 2025; Jiang, Liang, et al., 2025). This study provides demographic estimates of queer people as not only individuals, partners, or parents but also as siblings, cousins, and children, adding quantitative estimates to a body of mostly qualitative literature (e.g., Robinson et al., 2023; Stone et al., 2022). In doing so, this study offers broader implications for the social acceptance of queer people, namely by describing the demographic context to queer kinship.

Background

Toward a Queer Kinship Demography

Queer family demography raises fundamental questions about who is counted, and how, in demographic research (Fish et al., 2024; Moore & Stambolis-Ruhstorfer, 2013). Measurement issues, namely how to identify queer people and their children in large-scale surveys, are thus a

major focus of this literature (e.g., Baumle & Compton, 2014; Julian, Tessler, et al., 2025). Other work focuses largely on solidarity and estrangement within queer people's nuclear family (Bosley-Smith & Reczek, 2024; Fischer & Kalmijn, 2021; Reczek, 2025; Reczek et al., 2023; Weston, 1997). Parallel to queer family demography is the demography of kinship literature, which often assumes the primacy of biological kin ties and the universality of heterosexual partnership and procreation (Weston, 1997). Despite calls for research on understudied kinship patterns, including those of sexual minority families (Alburez-Gutierrez et al., 2022; Furstenberg et al., 2020; Weston, 1997), these literatures have lacked a demographic lens on queer kinship, leading to a gap regarding our understanding of the demography of queer kin.

A queer kinship demography, by contrast, critically re-evaluates the heterosexual assumptions of demographic kinship modeling (Allen & Mendez, 2018). The focus of study then shifts from an individual- or couple-level understanding of queer identity to one that recognizes queer people as members of nuclear and extended family networks. This reframing illuminates how population trends in queer identification are experienced at the individual level, with queer people filling the ranks of sibling, cousin, or aunt, among other kin types. To our knowledge, however, there are no estimates of how many people have queer relatives, let alone how queer identity is distributed within kinship networks.

Such an approach is crucial because queer people have historically been denied kinship both socially and legally (Weston, 1997). Though acceptance of queer sexual identities has grown, non-acceptance from family members means many queer individuals continue to face strained or even estranged family relationships (Reczek, 2016; Reczek et al., 2023). Equal family recognition under the law has also not been afforded to queer people; for example, even after marriage equality was recognized nationally in 2015 in the United States, queer people still struggle to become parents let alone have their parental rights recognized (Baumle & Compton, 2014; Gash & Raitskin, 2018; Park et al., 2016). This history of marginalization from kinship motivates the need to focus on biological kin who are, by definition, the family we do not choose. In doing so, we emphasize that queer people do not exist outside of traditional family structures but are instead integral, and often sizeable, parts of them.

Queer Sexual Identification in the United States

Queer people represent a significant population in the United States, at nearly nine percent of

adults (Gallup, 2025; Julian, Manning, & Westrick-Payne, 2024). As Panel A of Figure 1 shows, roughly one-third of young women identify as queer, compared to roughly one-sixth of young men. At the highest, over 37% of 18-year-old women and 17% of 19-year-old men identify as queer. This share then declines with age, with women and men reporting a queer identity at rates of 4.9% and 6% at age 50, respectively. Notably, by the early 50s, the trend shifts: from age 52 and onward, more men than women report a queer identity, but the difference between genders is rather small—at most, 58-year-old men are 2.3 percentage points more likely than women to identify as queer.

Panel B of Figure 1 then disaggregates these patterns into bisexual, gay/lesbian, and other sexual identity. Here, it is evident that the vast majority of young women who identify as queer identify as bisexual relative to lesbian or another identity. By contrast, a greater share of queer-identifying men identifies as gay, though a substantial share of young men do identify as bisexual. These patterns reflect established cohort and period trends in sexual identification (Julian, Manning, & Westrick-Payne, 2024; Martin & D’Augelli, 2007; Silva, 2025).

[INSERT FIGURE 1 ABOUT HERE]

Because individuals are embedded in larger extended kinship networks that change in composition over the life course as kin are born and die (Caswell, 2019; Jiang, Zuo, et al., 2025), the age and sex patterns of queer identification presented in Figure 1 thus suggest large cohort differences in the experience of having queer kin. Moreover, even a low share of the population identifying as queer implies a non-trivial number of people will have at least one queer kin. Yet, how queerness is distributed across the kinship network—in terms of the age, sex, and relationship of queer kin—is unexplored.

Current Study

This study estimates kin counts, including the expected number of kin who identify their sexuality as queer (i.e., lesbian/gay, bisexual, or other non-straight identity), among other identities. Using a mathematical demographic model that estimates the expected number of living kin—including kin type, age, and gender—for an average person who we call “Focal” (Caswell, 2022; Caswell & Song, 2021) and U.S. survey data on queer identification, we generate period-specific population-representative estimates of the experience of queer kinship for average people from different U.S.

birth cohorts. We focus on the horizontal and downward kin of Focal—siblings, nibblings, aunts/uncles, cousins, children, grandchildren—and exclude parents, grandparents, and great-grandparents because these ancestors may have had different fertility patterns depending on their identity, and our method is not suited for disentangling these historical patterns. Further, we describe heterogeneity in the age and sex composition of queer kin. We then produce estimates of the share of the U.S. population with at least one queer kin and describe differences by genealogical closeness, that is, whether they are nuclear or extended kin.

Of note, it is likely that many people are unaware they have any queer kin. Sexual identity disclosure is selective, shaped by genealogical distance, contact frequency, and the navigation of risks regarding unwanted outing (Grafsky et al., 2018; Scherrer, 2016; Weston, 1997). Bisexuals are especially unlikely to disclose their sexual identity to family members, particularly when in a different-sex partnership (Doan & Mize, 2020; Scherrer, 2016; Scherrer et al., 2015). Our intent therefore is not to describe family dynamics amongst kinship networks with an “out” member. Rather, we describe how a variety of demographic and social factors—including age and sex patterns of queer identification and kinship network age and sex structure—contribute to the number of queer kin an average person in the United States has, whether or not they are aware of their kin’s sexual identity.

Data & Methods

To estimate the number of queer kin, we combine mathematical demographic methods with sociodemographic and survey data. Estimates are produced for individuals aged 0-100 under the average kinship structure projected for 2024 and by kin type, sex, and age. Our empirical approach includes several steps: (1) calculate the number of living kin (e.g., siblings) for Focal; (2) estimate how many of these kin are queer and how this varies by kin age and sex; and (3) calculate the population share with one or more queer kin. For simplicity, we present results for a female Focal; results for a male Focal are substantively similar.

Demographic rates: mortality and fertility schedules

Period-, age-, and sex-specific mortality and fertility schedules come from the Human Mortality Database (2026), Human Fertility Database (2026), and Human Fertility Collection. We use the 1933-2024 data for the first two sources, which represent all available data at this time. We follow

Abrams et al. (2025) to impute male fertility using the 1969-2015 rates provided in the Human Fertility Collection; see Appendix A for more information on the approach.

Queer identification rates: queer schedule

Data on sexual identification are drawn from the Phase 3.2–Phase 4.2 releases of the Household Pulse Survey (HPS), a cross-sectional dataset fielded between July 2021 and September 2024 by the U.S. Census Bureau. Respondents aged 18 and older report on their sociodemographic characteristics including sexual identity, and how various social and economic issues (e.g., COVID-19) impacted their households. The HPS provides population estimates of queer identity comparable to similar surveys, with the advantage of being nationally representative (Brady et al., 2025; Deng & Watson, 2023; Julian, Manning, & Kamp Dush, 2024).

Comparing the HPS to similar data such as the National Health Interview Survey (NHIS), General Social Survey (GSS), or Behavioral Risk Factor Surveillance System (BRFSS) provides confidence in this choice (see Appendix B), namely because the HPS is fielded in all 50 states and has better coverage of older Americans, who are often underrepresented in surveys (Fredriksen-Goldsen & Kim, 2015). The queer schedules in HPS are most similar to those in BRFSS, though BRFSS slightly underestimates the share of queer young people, likely due to differences in interview format (Deng & Watson, 2023; Julian, Manning, & Westrick-Payne, 2024) because queer people may be hesitant to disclose their sexual identity in phone or face-to-face surveys compared to online (Lee et al., 2018). We return to these comparisons in the Results section.

The HPS was designed with low response rates in mind, as its goal is to quickly collect high frequency data on the “pulse” of American households. All estimates are survey weighted to ensure they are nationally representative and we calculate 95% confidence intervals using survey-weighted standard errors derived from the 80 replicate weights provided by the HPS. Variance estimation follows the Census Bureau’s successive difference replication method, which accounts for complex survey design and weighting adjustments, thereby tempering the potential impact of low response rates.

From these data we estimate a queer schedule, which describes age and sex-specific rates of queer identification. Individuals are coded as *queer* if they report identifying as gay, lesbian, bisexual, or “something else/other” and *heterosexual* if they report identifying as straight. Respondents who report identifying as “questioning” or “unsure” are coded as *unsure*, and

respondents who did not respond to the question are coded as *unknown*. Lacking identification data on individuals younger than 18 or older than 87 (age is top coded at 88), we code these ages as *unknown*. More information on this measure can be found in Table S1 in Appendix B.

The decision to include individuals who report their sexual identity as “something else/other” as queer follows the literature (e.g., Cheng, 2026; Mittleman, 2023). This measurement is more inclusive of the queer experience given it provides a response option for individuals who do not identify as lesbian, gay, or bisexual and do not use the label of heterosexual, which is more common among younger cohorts in the U.S. (Julian, Manning, & Westrick-Payne, 2024; Silva, 2025). Still, this label is arguably less coherent than the clearer LGB identities captured under the broader queer umbrella. As such, we estimate supplementary models in which these individuals are recoded as “straight,” which results in minimal changes to our findings given the relatively small share of individuals who report their identity thusly; we describe this further in the Results.

Though queer sexual and gender identities are often intertwined, our analytic focus is queer sexual identity. This focus is for the most part technical given our demographic models are two-sex and parameterized using sex-specific fertility and mortality schedules, and non-cisgender populations are not measured in the demographic data. Doing so allows us to align model assumptions with observed demographic rates and to interpret sex as a proxy for gender. We emphasize this exclusion reflects data and modeling constraints rather than a conceptual separation of sexual and gender identity, and we return to this limitation in the Discussion.

Analytic strategy

Our analysis proceeds in several steps. In the first, we apply the two-sex, time-variant kinship matrix model (Caswell, 2022; Caswell & Song, 2021) using the R package DemoKin (Williams & Alburez-Gutierrez, 2023) to generate estimates of the expected number of aunts/uncles, cousins, siblings, niblings (nieces and nephews), children, and grandchildren for an average person (henceforth “Focal”) aged 0-100.¹

A matrix formulation of the cohort-component method, the kinship model uses sex, age, and period-specific mortality and fertility rates to iteratively generate kin counts, including the age

¹ We also consider including adolescents aged 14-17 surveyed in the Youth Risk Behavior Survey (YRBS) but do not present these results due to discontinuities in the age-sex patterns of queer identification between data sources. Findings were substantively similar and can be found in Appendix C.

and sex composition of the kinship network, for a Focal aged between 0-100, for each year 1933-2024, and across kin type. In each year, kin counts are determined by (1) the corresponding fertility rates to birth new kin, and (2) the corresponding survival ratios to age and die kin. Put simply, the expected number of kin at age x of Focal depends on the number alive at age $x-1$ who survived over the year interval, plus new kin born during this interval. Average kinship structures estimated from these models are comparable against empirical data (Alburez-Gutierrez et al., 2023; Calderón-Bernal et al., 2025; Van Damme et al., 2025). This modeling strategy follows from previous work that has been used to calculate prevalences of kin characteristics such as bereavement due to war (Acosta et al., 2026), unemployment (Song & Caswell, 2022), dementia (Feng et al., 2024), and educational attainment (Adhikari & Alburez-Gutierrez, 2025; Jiang, Liang, et al., 2025).

Notably, sexual identity can be fluid, changing across the life course (Mittleman, 2023; Saperstein, 2025). As our data are cross-sectional, we do not use multistate kinship models (Caswell, 2020) because they require age- and sex-specific probabilities that an individual changes sexual identity between ages. Such an approach is beyond the scope of this study, as we are primarily interested in producing a representative snapshot of queer kinship in the United States. This notwithstanding, fluid sexual identity, which is most prevalent among younger cohorts of women, changes most often between the straight, bisexual, and “something else” categories (Mittleman, 2023). We therefore do not expect fluidity to significantly alter our results for most categories (i.e., for kin who consistently identify as gay or lesbian) or for older kin. The transitions between a straight and queer (e.g., bisexual) identity, then, may be less problematic than other sexual identity transitions; to remedy this, we present our period-specific results disaggregated by sexual identity. We return to this in the Discussion.

The second step of the analysis involves applying the kinship estimates from the first step to the queer schedule in a manner akin to previous work using the kinship matrix models (see equation 1; Caswell, 2019; Feng et al., 2024; Song & Caswell, 2022). This generates estimates of the expected number of queer kin among the aunts/uncles, cousins, siblings, niblings, children, and grandchildren of Focal. We first present estimates for all kin and then by the age and sex of kin. These estimates represent population-level averages, assuming the projected 2024 kinship structure and pooled rates of queer identification.

Let the number of living kin aged α , sex γ , and type κ at Focal age χ be represented by the

vector $N_{\chi,\alpha\gamma\kappa}$, and the age-sex-specific prevalence of queer identification the vector $R_{\alpha\gamma}$. We obtain the expected number of queer kin by multiplying these two vectors:

$$Q_{\chi,\alpha\gamma\kappa} = N_{\chi,\alpha\gamma\kappa} \cdot R_{\alpha\gamma}, \quad (1)$$

where $Q_{\chi,\alpha\gamma\kappa}$ represents the expected number of queer kin of age α , sex γ , and kin type κ for a Focal aged χ . Note that this estimate represents a snapshot of queer kin under 2024 rates; we do not assume the queer prevalences are fixed and we discuss this assumption in greater detail in the Discussion section. We sum over all α and γ to produce the expected number of queer kin of type κ at Focal age $\chi \in (15, 30, 45, 60, 75)$ in 2024:

$$Q_{\chi,\kappa} = \sum_{\alpha\gamma} Q_{\chi,\alpha\gamma\kappa}. \quad (2)$$

In the third step, we follow Acosta et al. (2026) in using the individual-level estimates of queer sexual identification to calculate the share of U.S. residents with one or more queer kin, distinguishing between three categories of kin relationship represented by ζ : any kin type, nuclear kin (siblings and children), and extended kin (aunts/uncles, cousins, niblings, and grandchildren). Note that these estimates exclude spouses, step-kin, and other non-biological family. We first calculate the age- and sex-specific probability to have queer kin of type ζ using its complement, the probability that Focal aged χ has no queer kin of type ζ :

$$p_{\chi,\zeta}^0 = \prod_{\alpha} \prod_{\gamma} (1 - q_{\gamma\alpha})^{N_{\zeta\gamma\alpha\chi}} \quad (3)$$

where $q_{\gamma\alpha}$ is the age- and sex-specific probability to be queer, and $N_{\zeta\gamma\alpha\chi}$ is the average number of living kin. Finally, we estimate the share of people who have at least one queer kin by taking the complement of the probability that a Focal of a given age has no queer kin of a given type. This formula is written as:

$$f_{\zeta\theta}^A = (1 - p_{\chi,\zeta}^0). \quad (4)$$

We check the robustness of this measure by instead assuming that the number of queer kin follows a Poisson distribution (Caswell, 2024); doing so yields substantively similar results (see Table S5). On average, these two estimates differ by one percentage point, with a maximum of four percentage points of absolute difference.

Assumptions of approach

This modeling strategy rests on several assumptions. First, we assume that the age and sex distributions of queer identity within a kinship network follows the age and sex distributions of

queer identity within the population, and that the queer schedule reflects these distributions. Second, we do not assume an intergenerational transmission of queerness (e.g., that children of same- or different-sex parents are more likely to be queer or heterosexual, respectively) or that sexual identity clusters within families. We return to the implications of these assumptions in the Discussion section.

Additionally, the fertility and mortality schedules apply to queer and non-queer populations alike, which we argue should not meaningfully bias our estimates. Given that queer identity is mutable, constructing separate demographic schedules would require making assumptions about queer identification across the life course that are beyond the scope of this study. Moreover, our approach does not rely on knowing queer-specific fertility, such as the number of grandchildren we have through our queer children but rather the number of queer children and grandchildren. Differential mortality, however, poses a more substantive concern, as it affects kin survival regardless of generation.² Insofar as a mortality disparity is real and stable—many studies proxy queer identity through sexual behavior, which can generate different findings than direct measures of identity (Matthews et al., 2014)—our model would somewhat overstate the number of living queer female kin, particularly for older cohorts of kin with the highest cumulative mortality exposures. Additionally, HIV-related mortality, which disproportionately affected queer men in the 1980s and 1990s, is already captured in the mortality schedules for these periods; consequently, estimates for cohorts exposed to this risk should already reflect excess mortality among queer men. To the extent this affects our findings, we anticipate it would most affect uncles who lived through the HIV epidemic and were sexually active with men; moreover, HIV risk was primarily contingent on sexual behavior rather than identity.

Results

Sexual identity within the kinship network

Figure 2 shows the sexual identity composition of living kin disaggregated by kin type for an average (Focal) person aged 15, 30, 45, 60, and 75 in 2024 (see Table S3 for the corresponding values, including standard errors). The size and composition of Focal's kinship network vary

² Evidence is mixed, with some studies suggesting that queer women's suicide risk is higher than heterosexual women's, though suicide was a small share of all deaths. These studies also find few differences in male mortality outside of HIV-related causes (Cochran & Mays, 2015; Laughney & Eliason, 2022).

greatly by age due to different mortality and fertility trends experienced by cohorts throughout their lifetimes. For example, a 30-year-old born in 1994 is expected to have more living aunts and uncles than a 15-year-old born in 2009—not because of greater mortality experienced by the latter person’s aunts and uncles, but because of higher and earlier fertility of the former’s grandparents.

[INSERT FIGURE 2 ABOUT HERE]

Under the 2021-2024 rates of queer identification, Focal can expect important variation in both the number and composition of queer kin. For example, though Focals born in 1949 and 1964 can expect to have similar numbers of living children, a greater share of the younger person’s children is expected to be queer, reflecting the younger ages of their children. Similarly, though Focals born in 1979 and 1994 are expected to have similar numbers of cousins, a smaller share of the latter person’s cousins is expected to be heterosexual—even though a greater share has “unknown” sexual identity. This finding is due both to the age patterning of queer identification and age composition of kin, as the younger person’s cousins are predominantly younger than 35, meaning they are more likely to report a queer identity or be classified as “unknown.”

Age of queer kin and queer peers

Next, we explore how age-specific queer identification patterns are reflected in kinship networks by examining the age of Focal simultaneously with the age of their queer kin. We focus on aunts/uncles, siblings, and cousins because previous research has identified them as particularly supportive regardless of if Focal or their kin is queer (Grafsky et al., 2018; Reczek et al., 2022; Stone et al., 2022). Each panel in Figure 3 displays a normalized expected number of queer kin, with lighter yellow shades depicting the Focal–kin age combinations with a relatively higher number of queer kin, and darker blue shades depicting Focal–kin age combinations with a relatively lower number of queer kin. The 99th, 75th, and 50th percentiles of these values are shown by the solid, dashed, and dotted lines, respectively, and are specific to each panel.

The solid line in panel “All kin” suggests there is one age range where Focal has the greatest number of queer kin of any type: very early adulthood, and the remaining panels confirm this is largely due to queer cousins and siblings. The 99th percentile line (solid) suggests that Focals with the highest queer kin counts are young, with similarly young kin. The lines for the 75th and 50th percentiles by contrast indicate that whereas individuals of all ages may have queer kin, the

counts are especially likely when kin are under age 40 (dashed line) or when both Focal and kin are older (dotted line).

Focals with the greatest expected number of queer aunts/uncles tend to be infants (panel “Aunts/Uncles”), whereas Focals with the highest quantity of queer siblings and cousins tend to be in their late teens and early twenties (panels “Siblings” and “Cousins,” respectively). In these three panels, the 99th percentile line is always drawn around young Focals whose kin are likewise young. Panel “Aunts/Uncles” shows that queer aunts/uncles are less common for older Focals. Panels “Siblings” and “Cousins” instead suggest that queer kin in these categories are similarly aged to Focal. Together, these figures show that younger people have more queer kin because the bulk of their living kin are also young.

[INSERT FIGURE 3 ABOUT HERE]

We also define and describe *queer peers*, that is, horizontal queer kin (i.e., not direct ancestors or descendants) with an age difference amenable to creating a positive, peer-like relationship. For siblings and cousins, this includes kin within five years older or younger (e.g., for a 20-year-old Focal, their queer kin are 15-25). Because aunts and uncles may be especially supportive when they can intercede in the parent-child relationship, they need to be appropriately older than Focal (Grafsky et al., 2018), so we set this age difference as 10 years and older.

Figure 4 shows important variation in the number of queer peers across Focal age and kin type. Queer peer aunts and uncles decline in number as Focal ages. For siblings and cousins, however, the number of queer peers peaks during the early twenties before declining steadily thereafter. This pattern reflects two competing dynamics: Focals under age 22 have a limited window of queer peers because the lower bound of the queer schedule is 18, but as Focal ages into the early twenties, more of their peer kin fall within the age range with the highest rates of queer identification, driving the initial increase. The subsequent decline reflects the age gradient of queer identification whereby as Focal ages past their mid-twenties, their queer peers are increasingly drawn from older ages where queer identification is lower which reduces the expected count of queer peers. These findings confirm that younger Focals’ queer kin are often of the same generation compared to older Focals, whose queer kin are primarily descendants.

[INSERT FIGURE 4 ABOUT HERE]

Male and female queer kin

Finally, we describe how sex differences in queer sexual identity contribute to large cohort differences in the sex composition of Focal's kinship structure (Figure 5; see also Figure S11 and Table S4 in the Online Appendix). Regardless of cohort, Focal has a larger number of queer female kin than queer male kin. Moreover, the age pattern of queer identification drives sex differences in queer kin: younger Focals have more queer female horizontal kin (siblings, cousins), whereas older Focals have more queer female descendants (children, niblings, grandchildren). Though more men than women report being queer from the early 50s onward, these sex differences in queer kin counts can be attributed to the fact that drastically more women than men of younger ages are queer.

[INSERT FIGURE 5 ABOUT HERE]

Population share with queer kin

We then turn from an individual-level view of the experience of having queer kin to describe the share of the population with at least one queer kin. Note that the columns are not mutually exclusive, meaning individuals could have queer close and extended kin, and neither is reflected in the calculation for the other.

As shown in Table 1, the share of the population with queer kin of any age and kin type is large, ranging from a low of 74.8% among 15-year-olds to a high of 88.2% among 45-year-olds. That most individuals of all ages can expect to have at least one queer family member is driven largely by the higher prevalence of bisexual identification relative to gay/lesbian or "other" identification, especially at younger ages (see Figure S13). The age patterning of bisexual identification and the prevalence of young extended kin jointly shape our finding that individuals are substantially more likely to have at least one bisexual relative compared to a relative who identifies as gay/lesbian or another non-straight identity.

Although a large share of U.S. residents has one or more queer kin, this prevalence is driven overwhelmingly by queer relatives in the extended family, particularly bisexual extended kin. When disaggregated by genealogical closeness, people are far less likely to have queer nuclear kin, including siblings and children, than extended kin (e.g., niblings, grandchildren). For example, among individuals aged 30, 29.1% have at least one queer nuclear kin, while 82.3% have at least

one queer extended kin. These estimates suggest that a substantial share of the U.S. population has queer relatives in their nuclear and/or extended families.

[INSERT TABLE 1 ABOUT HERE]

Robustness checks and supplementary models

Because estimates of queer sexual identity varies—often greatly—between surveys and across time (Deng & Watson, 2023; Julian, Manning, & Westrick-Payne, 2024), we conduct a cross-period comparison of the HPS queer schedule with three other sources: the NHIS, GSS, and BRFSS. Appendix B describes the differences in more detail; in short, we find the HPS advantageous over other sources because participants make up a wider age range across all U.S. states, making it more representative of the U.S. population. Although our goal is to provide a cross-sectional snapshot of queer kin in the U.S., we argue our findings demonstrate remarkable concordance with those produced using similar surveys, and variation may be attributable to differences in survey mode (i.e., face-to-face, phone, or online; Deng & Watson, 2023). Using the projected 2024 kinship structure and pooled 2021-2024 rates of queer identification for the four sources, we estimate the number of queer kin as a share of total kin. These estimates range from 6.6% in the NHIS to 12.5% in the HPS, with 10.3% in the BRFSS (Figure S6 in Appendix B), and differences between surveys arise from variation in estimates among young people.

Beyond validating our choice of data source, this comparison also functions as a sensitivity analysis with respect to the queer schedule. Because the four sources span different collection windows, Figure S6 quantifies how our estimates would differ under earlier rates. For example, the queer share of an average 30-year-old Focal’s kinship network more than tripled between 2012 and 2024, reflecting the substantial increase in queer identification over this period. Our estimates should therefore be read as period-specific descriptions of queer kinship under contemporary rates rather than as reflecting stable levels; if identification rates change, the composition of kinship networks will shift accordingly.

Finally, we assess robustness of the results based on the coding of sexual identity by replicating the analyses by reclassifying individuals who report their identity as “something else/other” so the measure of queer identity includes only lesbians, gays, and bisexuals (Cheng, 2026). Doing so affects female kin more than male kin, as up to six percent of women aged 18-21 in the United States identify as “something else/other” compared to three percent of similarly aged

men. Specifically, this reclassification decreased the expected number of queer kin by about 17 percent, or roughly 0.3 total kin. However, we do not present these estimates as our main result because the queer demography literature recognizes them as part of the broader queer umbrella (e.g., Cheng, 2026; Mittleman, 2023), and younger cohorts increasingly rely on these labels to describe their sexual identity (Julian, Manning, & Westrick-Payne, 2024; Silva, 2025).

Discussion

This study advances a queer kinship demography that focuses on queer people as members of extended kinship networks, emphasizing queer people as part of, rather than external to, family life (Weston, 1997). We focus on biological kin to emphasize how demographic processes interact with social phenomena (e.g., age and sex patterning of queer identification) to shape exposure to queer relatives within the extended family network. Combining demographic kinship modeling techniques with survey data that directly identifies sexual minorities, we produce novel period estimates of queer kin in the United States. At the individual level, we estimate how many queer kin individuals have across birth cohort and describe age and sex dynamics that contribute to these counts; at the population level, we estimate the share of individuals with at least one queer kin and describe differences by kin sexual identity and genealogical closeness.

Our results demonstrate that queer kin counts reflect the interaction of several demographic processes: age- and sex-specific patterns of queer identification, kinship age and sex structure, and the historical fertility and mortality regimes that shape kinship network size and composition as well as population age and sex structure. We find substantial cohort differences in the composition of queer kin: younger cohorts are more likely to have queer aunts and uncles, cousins, and siblings, whereas older cohorts are more likely to have queer niblings, children, and grandchildren. These findings are consistent with prior surveys showing that younger individuals are more likely than older individuals to report having a gay or lesbian close friend or family member (Becker, 2012; Neidorf & Morin, 2007). Moreover, we find that over 75 percent of the U.S. has at least one queer relative, driven largely by (bisexual) extended kin who are more numerous than nuclear kin (Leopold et al., 2025).

Our effort to provide estimates of queer kinship offers two main contributions to the literature. First, individuals experience population trends through kinship networks. The consequences of delayed and declining fertility and mortality in the United States are felt through

smaller kinship structures (Alburez-Gutierrez et al., 2023; Caswell, 2019) which are shifting toward more of a “beanpole” shape, with the prevalence of horizontal kin (i.e., siblings, cousins) overtaken by increasingly overlapping lifetimes with vertical kin (i.e., grandparents, grandchildren) (Bengtson, 2001; Jiang, Zuo, et al., 2025; Verdery, 2015). In estimating period-specific counts of queer biological kin (i.e., excluding partners, in-laws, step-kin, and “chosen family”), we expand the understanding of queer kinship from a focus on nuclear family members to one situated within a shifting demographic context and emphasize one way that demographic change is felt at the family level—that is, through the experience of having queer kin.

Second, we extend the queer family demography literature—which has focused more narrowly on same-sex partnership and parenthood (Allen & Mendez, 2018)—to focus on a broader range of nuclear and extended kin, many of whom are understudied: aunts/uncles, cousins, siblings, nieces, children, and grandchildren. In doing so, our study provides demographic context for qualitative work, for example findings that highlight the importance of extended family ties for young queer people, ranging from the provision of housing and material resources to assistance in navigating complicated family dynamics (Reczek et al., 2022; Robinson et al., 2023; Scherrer, 2016; Stone et al., 2022). These ties can be especially consequential when extended kin are themselves queer: for example, queer aunts, uncles, and cousins can serve as role models or safe harbors for youth navigating family heteronormativity (Robinson et al., 2023; Stone et al., 2022). Understanding queerness as constituent of family networks rather than as external to family life thus has direct implications for queer people’s access to support and kinship. Though queer people are often denied full participation in kinship either socially or legally (Weston, 1997), our findings underscore the near ubiquity of queer kinship ties in the United States and highlight the importance of expanding queer kinship beyond the nuclear family.

Third, we develop a novel technique for estimating queer kinship using relatively simple demographic and survey data as inputs. Our approach offers multiple dimensions through which to study queer kinship, including the demographic characteristics of both Focal and their kin, and the population-level exposure to queer kin in both the nuclear and extended family. This technique can be further extended to describe heterogeneity in queer kinship between groups (e.g., race-ethnicity). Moreover, we provide confidence in our estimates through comparisons with similar surveys. We conclude our estimates provide an upper bound to queer kinship in the United States

which may reflect differences in survey mode; sexual minorities likely feel more comfortable disclosing their identity in online surveys versus face-to-face or by phone (Deng & Watson, 2023).

Of note, individuals may be unaware they have queer relatives. For many sexual minorities, queer identity disclosure to family members occurs after the formation of a same-sex relationship, which can delay or limit family awareness of their sexual identity. This dynamic can be especially pronounced among bisexuals, who may be perceived as heterosexual when in different-sex partnerships (Doan & Mize, 2020; Scherrer, 2016; Scherrer et al., 2015). Queer individuals may also be less likely to disclose their identity to extended kin compared to their nuclear family, as individuals tend to be closer with the latter (Leopold et al., 2025). Thus, the *true* prevalence of queer kin within family networks is higher than the *known*; this distinction matters because it translates to decidedly different realities for queer people's family relationships. If the gap between the true and known prevalence is large, such as in more conservative contexts, social acceptance of sexual minorities will be lower than in contexts where the gap is small and people are aware of their queer family members (Becker, 2012). There may be important cohort differences in this prevalence gap, with the gap likely converging over time and by cohort as acceptance of sexual minorities increases.

For this reason, our approach is advantageous over respondent reports of the sexual identity of their kin—where available, these data are often unrepresentative and likely to dramatically underestimate the sexual identity composition of individuals' kinship networks. For example, a module in the 2006 General Social Survey asked respondents how many of their family members were gay men or women, but the sample was small ($N = 491$) and is now two decades out of date. Other surveys conflate close friends with family, making it impossible to disaggregate the two. Examining rates of sexual identity disclosure within the kinship network and how these patterns shape the visibility of queer family members is thus a useful avenue for future research.

We assume that queer sexual identity does not cluster within families. In this vein, we ignore potential relationships between the family or social environment that can shape sexual identity (e.g., conservative attitudes that may drive both higher fertility and lower support of sexual minorities) for both queer and heterosexual individuals (Silva, 2019; Silva & Evans, 2022). In our approach, a large family would have the same distribution of queer kin as a small family; in turn, the larger kinship network would have more queer family members. Previous investigations into such clustering relied primarily on small convenience samples with limited population

representation (e.g., Bailey & Benishay, 1993; Kendler et al., 2000; Pillard, 1986). To the extent that queer identity clusters within families, we may underestimate queer kin counts in some families but overestimate them for others, though population-level averages should be unaffected.

This study has some limitations. First, we focus on queer sexual identity to the exclusion of queer gender identity. The primary reason is technical: the kinship models and demographic input data assume the mortality and fertility patterns of cisgender people, and we describe above why these patterns can be assumed for queer and non-queer cisgender people alike. Substantively, our goal is to understand how age and sex patterns of queer identification interact with kinship structure to shape experiences of having queer kin by age group. Producing estimates for gender diverse populations would be a useful follow-up to this study.

Second, measuring sexual identity is complex. Identification varies by age group, with younger cohorts far more likely to report bisexual or “something else” for sexual identity (Julian, Manning, & Westrick-Payne, 2024). Our inclusive definition of queer identity likely minimizes the impact of this limitation, but our relatively limited age range (18-87) reflects a common challenge in studying queer populations at the extremes of the age distribution (Fredriksen-Goldsen & Kim, 2015; Macapagal et al., 2017). In a similar vein, sexual identity can be fluid across the life course (Julian, Newmyer, Manning, et al., 2024; Mittleman, 2023; Saperstein, 2025). Although addressing fluidity is beyond the substantive and methodological scope of this study, we present estimates disaggregated by sexual identity label which suggest our results are driven by a larger share of kin identifying as bisexual (see Table 1 and Figure S13). Because prior research has demonstrated that most sexual identity changes are between the straight, bisexual, and “something else” categories (Mittleman, 2023), we can assume that our results for gay and lesbian kin are most stable. Moreover, because young women are overrepresented in the sexually fluid population, we expect that Focals with a greater number of young female kin, such as 15- and 30-year-olds, would similarly have a greater number of kin who ever identified as queer. Including sexual fluidity through multistate models would be a useful extension of this paper, given a significant share of individuals may change between queer and straight identities over time.

Third, the kinship matrix model defines kinship as biological, to the exclusion of adopted kin, in-laws, and “chosen family.” As family is “something we *do*, rather than something we *have*” (Allen & Mendez, 2018, p. 71), this is an important limitation. In our view, ours is a first attempt at “queering” kinship demography, meaning a critical re-evaluation of (cis)heterosexual

assumptions in demographic modeling such as the universality of heterosexual partnership and procreation (Allen & Mendez, 2018; Weston, 1997). The study contributes empirical estimates on queer kin in a case where no comparable empirical data exist. Whereas household-level estimates of queer individuals, couples, and parents are increasingly available (e.g., Baumle & Compton, 2014; Julian, Westrick-Payne, et al., 2025), information on queer kin outside the nuclear family or household remains scarce. Our estimates therefore do not reflect the full expanse of queer kinship in the United States when full kinship ties—legal, marital, or chosen—are included. Developing models and collecting data that incorporate non-biological and chosen kin remain important directions for research in this area.

Fourth, we do not model subnational heterogeneity by race-ethnicity, region, education level, nativity, or other relevant variables. We replicate Figure 1 stratified by these characteristics in Appendix E and conclude that subgroups have similar identification patterns, suggesting that subnational heterogeneities in queer kinship would be due to differences in kinship size and structure rather than differences in queer identity. Focusing on heterogeneity in queer kinship across a number of dimensions is a promising direction for future work, however.

Finally, our estimates are period-specific, raising the question of how these patterns could change over time given age, period, and cohort dynamics of queer identification are not well understood (Julian, Manning, & Westrick-Payne, 2024; Martin & D’Augelli, 2007). Changes to age and sex patterns of queer identification, in addition to continued changing demographic processes, will be reflected in different distributions of queerness within kinship networks. Continued improvements in the measurement of sexual and gender identity are critical for understanding how these identification dynamics interact with kinship structure.

Contact theory suggests that interacting with minority group members (e.g., sexual minorities) increases tolerance among majority group members (e.g., heterosexuals) (Allport, 1954). Because many people have homophilous social networks, meaningful contact with queer people for many heterosexuals may be limited to their family ties (McPherson et al., 2001). Such contact is likely not rare—though fewer than 10 percent of Americans identify as queer, at least 75 percent of Americans have a queer relative. This study thus has important implications for understanding broader patterns of queer identification and the potential downstream effects on social acceptance given knowing queer people, even as acquaintances, can increase support of LGBTQ+ people and their rights (DellaPosta, 2018; Rosenfeld, 2017). For example, these findings

may help explain why young people tend to espouse greater support for LGBTQ+ issues (Becker, 2012), because they are more likely to be exposed to queer peers in both friendship and family networks. By extension, increasing public knowledge of how queer identity is distributed across the family and how many people have queer kin, could be vital for increasing public support for sexual minorities and improving the family relationships of queer people. Through documenting the prevalence of queer kinship in the United States, our results identify a demographic mechanism linking exposure to queer people to broader social trends.

REFERENCES

- Abrams, L., Jiang, S., Alburez-Gutierrez, D., Acosta, E., & Zagheni, E. (2025). *Adverse Mortality Trends have Increased the Number Bereaved in the United States*. SocArXiv. https://doi.org/10.31235/osf.io/z38cs_v1
- Acosta, E., Alburez-Gutierrez, D., Gargiulo, M., & Torres, C. (2026). Weaponizing Kinship: A Demographic Analysis of Bereavement in the Colombian Conflict. *Population and Development Review*, n/a(n/a). <https://doi.org/10.1111/padr.70048>
- Adhikari, S., & Alburez-Gutierrez, D. (2025). *The future of grandparenthood in South Asia: The role of population aging and educational expansion* (WP-2025-003; 0 ed., p. WP-2025-003). Max Planck Institute for Demographic Research. <https://doi.org/10.4054/MPIDR-WP-2025-003>
- Alburez-Gutierrez, D., Barban, N., Caswell, H., Kolk, M., Margolis, R., Smith-Greenaway, E., Song, X., Verdery, A., & Zagheni, E. (2022). *Kinship, Demography, and Inequality: Review and Key Areas for Future Development*. SocArXiv. <https://doi.org/10.31235/osf.io/fk7x9>
- Alburez-Gutierrez, D., Williams, I., & Caswell, H. (2023). Projections of human kinship for all countries. *Proceedings of the National Academy of Sciences*, 120(52), e2315722120. <https://doi.org/10.1073/pnas.2315722120>
- Allen, S. H., & Mendez, S. N. (2018). Hegemonic Heteronormativity: Toward a New Era of Queer Family Theory. *Journal of Family Theory & Review*, 10(1), 70–86. <https://doi.org/10.1111/jftr.12241>
- Allport, G. W. (1954). *The nature of prejudice*. Addison-Wesley.
- Bailey, J. M., & Benishay, D. S. (1993). Familial aggregation of female sexual orientation. *American Journal of Psychiatry*, 150(2), 272–277. <https://doi.org/10.1176/ajp.150.2.272>
- Baumle, A. K., & Compton, D. R. (2014). Identity Versus Identification: How LGBTQ Parents Identify Their Children on Census Surveys: LGBTQ Parent-Child Identities and Identification. *Journal of Marriage and Family*, 76(1), 94–104. <https://doi.org/10.1111/jomf.12076>
- Becker, A. B. (2012). Determinants of Public Support for Same-Sex Marriage: Generational Cohorts, Social Contact, and Shifting Attitudes. *International Journal of Public Opinion Research*, 24(4), 524–533. <https://doi.org/10.1093/ijpor/eds002>
- Bengtson, V. L. (2001). Beyond the Nuclear Family: The Increasing Importance of Multigenerational Bonds. *Journal of Marriage and Family*, 63(1), 1–16. <https://doi.org/10.1111/j.1741-3737.2001.00001.x>
- Bosley-Smith, E., & Reczek, R. (2024). Why LGBTQ Adults Keep Ambivalent Ties with Parents: Theorizing “Solidarity Rationales.” *Social Problems*, 71(1), 220–236. <https://doi.org/10.1093/socpro/spac007>
- Brady, L. A., Julian, C. A., & Manning, W. D. (2025). Variation Between LGBT Estimates and State Policy Context. *Population Research and Policy Review*, 44(1), 11. <https://doi.org/10.1007/s11113-025-09938-2>

- Calderón-Bernal, L. P., Alburez-Gutierrez, D., & Zagheni, E. (2025). Analysing Biases in Genealogies Using Demographic Microsimulation. *European Journal of Population*, 41(1), 34. <https://doi.org/10.1007/s10680-025-09756-4>
- Caswell, H. (2019). The formal demography of kinship: A matrix formulation. *Demographic Research*, 41, 679–712. <https://doi.org/10.4054/DemRes.2019.41.24>
- Caswell, H. (2020). The formal demography of kinship II: Multistate models, parity, and sibship. *Demographic Research*, 42, 1097–1146. <https://doi.org/10.4054/DemRes.2020.42.38>
- Caswell, H. (2022). The formal demography of kinship IV: Two-sex models and their approximations. *Demographic Research*, 47, 359–396. <https://doi.org/10.4054/DemRes.2022.47.13>
- Caswell, H. (2024). The formal demography of kinship VI: Demographic stochasticity and variance in the kinship network. *Demographic Research*, 51, 1201–1256. <https://doi.org/10.4054/DemRes.2024.51.39>
- Caswell, H., & Song, X. (2021). The formal demography of kinship III: Kinship dynamics with time-varying demographic rates. *Demographic Research*, 45, 517–546. <https://doi.org/10.4054/DemRes.2021.45.16>
- Cheng, K. J. (2026). From the Rainbow to the Sunset: Anticipated Availability of Informal Care for Older Sexual Minorities in the United States. *Population Research and Policy Review*, 45(3), 24. <https://doi.org/10.1007/s11113-026-10008-4>
- Cochran, S. D., & Mays, V. M. (2015). Mortality Risks Among Persons Reporting Same-Sex Sexual Partners: Evidence From the 2008 General Social Survey—National Death Index Data Set. *American Journal of Public Health*, 105(2), 358–364. <https://doi.org/10.2105/AJPH.2014.301974>
- DellaPosta, D. (2018). Gay Acquaintanceship and Attitudes toward Homosexuality: A Conservative Test. *Socius: Sociological Research for a Dynamic World*, 4, 2378023118798959. <https://doi.org/10.1177/2378023118798959>
- Deng, B., & Watson, T. (2023). *LGBTQ+ Data Availability: What We Can Learn From Four Major Surveys*. Brookings Institute. https://www.brookings.edu/wp-content/uploads/2023/04/20230417_CCF_DengWatson_MAF-1.pdf
- Dixon, A. R. (2024). Empty chairs at the dinner table: Black-white disparities in exposure to household member deaths. *SSM - Population Health*, 27, 101704. <https://doi.org/10.1016/j.ssmph.2024.101704>
- Doan, L., & Mize, T. (2020). Sexual Identity Disclosure among Lesbian, Gay, and Bisexual Individuals. *Sociological Science*, 7, 504–527. <https://doi.org/10.15195/v7.a21>
- Feng, K., Song, X., & Caswell, H. (2024). Kinship And Care: Racial Disparities In Potential Dementia Caregiving In The U.S. From 2000 To 2060. *The Journals of Gerontology, Series A: Biological Sciences and Medical Sciences*, glae106. <https://doi.org/10.1093/gerona/glae106>

- Fischer, M. M., & Kalmijn, M. (2021). Do adult men and women in same-sex relationships have weaker ties to their parents? *Journal of Family Psychology*, 35(3), 288–298.
<https://doi.org/10.1037/fam0000696>
- Fish, J. N., Reczek, R., & Ezra, P. (2024). Defining and measuring family: Lessons learned from LGBTQ+ people and families. *Journal of Marriage and Family*, 86(5), 1450–1467.
<https://doi.org/10.1111/jomf.12987>
- Fredriksen-Goldsen, K. I., & Kim, H.-J. (2015). Count Me In: Response to Sexual Orientation Measures Among Older Adults. *Research on Aging*, 37(5), 464–480.
<https://doi.org/10.1177/0164027514542109>
- Furstenberg, F. F., Harris, L. E., Pesando, L. M., & Reed, M. N. (2020). Kinship Practices Among Alternative Family Forms in Western Industrialized Societies. *Journal of Marriage and Family*, 82(5), 1403–1430. <https://doi.org/10.1111/jomf.12712>
- Gallup. (2025). *LGBTQ+ Identification in U.S. Rises to 9.3%*. Gallup.
<https://news.gallup.com/poll/656708/lgbtq-identification-rises.aspx>
- Gash, A., & Raitskin, J. (2018). Parenting without Protection: How Legal Status Ambiguity Affects Lesbian and Gay Parenthood. *Law & Social Inquiry*, 43(01), 82–118.
<https://doi.org/10.1111/lsi.12233>
- Grafsky, E. L., Hickey, K., Nguyen, H. N., & Wall, J. D. (2018). Youth Disclosure of Sexual Orientation to Siblings and Extended Family. *Family Relations*, 67(1), 147–160.
<https://doi.org/10.1111/fare.12299>
- Hünteler, B. (2022). A new perspective on the generational structures of families – Generational placements over the life course. *Advances in Life Course Research*, 51, 100450.
<https://doi.org/10.1016/j.alcr.2021.100450>
- Jiang, S., Liang, H., Alburez-Gutierrez, D., & Zagheni, E. (2025). *Human capital investment helps mitigate family caregiving challenges in aging China* (WP-2025-021; 0 ed., p. WP-2025-021). Max Planck Institute for Demographic Research.
<https://doi.org/10.4054/MPIDR-WP-2025-021>
- Jiang, S., Zuo, W., Guo, Z., & Tuljapurkar, S. (2025). Changing Demographic Rates Reshape Kinship Networks. *Demography*, 62(3), 899–922. <https://doi.org/10.1215/00703370-11996578>
- Julian, C. A., Manning, W. D., & Kamp Dush, C. M. (2024). Measurement opportunities for studying sexual and gender diverse partnerships in population-based surveys. *Journal of Marriage and Family*, 86(5), 1410–1431. <https://doi.org/10.1111/jomf.12957>
- Julian, C. A., Manning, W. D., & Westrick-Payne, K. K. (2024). Responses to Sexual and Gender Identity Measures in Population-Level Data by Birth Cohort: A Research Note. *Demography*, 61(1), 15–30. <https://doi.org/10.1215/00703370-11164985>
- Julian, C. A., Newmyer, L., Manning, W. D., & Dush, C. M. K. (2024). Visualizing Multiple Sexual Identity Responses Using a Network Approach. *Socius*, 10, 23780231241279129. <https://doi.org/10.1177/23780231241279129>

- Julian, C. A., Tessler, H., Manning, W. D., VanBergen, A. M., & Kamp Dush, C. M. (2025). Half of the Picture: A Research Note on Measuring the Sexual Identity Composition of Couples. *Demography*, 62(4), 1185–1201. <https://doi.org/10.1215/00703370-12186740>
- Julian, C. A., Westrick-Payne, K., & Manning, W. (2025). *Number of Children Living in Households Headed by Same-Sex and Different-Sex Couples, 2023*. National Center for Family and Marriage Research. <https://doi.org/10.25035/ncfmr/fp-25-33>
- Kendler, K. S., Thornton, L. M., Gilman, S. E., & Kessler, R. C. (2000). Sexual Orientation in a U.S. National Sample of Twin and Nontwin Sibling Pairs. *American Journal of Psychiatry*, 157(11), 1843–1846. <https://doi.org/10.1176/appi.ajp.157.11.1843>
- Laughney, C. I., & Eliason, E. L. (2022). Mortality Disparities Among Sexual Minority Adults in the United States. *LGBT Health*, 9(1), 27–33. <https://doi.org/10.1089/lgbt.2020.0482>
- Lee, S., Fredriksen-Goldsen, K. I., McClain, C., Kim, H.-J., & Suzer-Gurtekin, Z. T. (2018). Are Sexual Minorities Less Likely to Participate in Surveys? An Examination of Proxy Nonresponse Measures and Associated Biases with Sexual Orientation in a Population-Based Health Survey. *Field Methods*, 30(3), 208–224. <https://doi.org/10.1177/1525822X18777736>
- Leopold, T., Raab, M., Becker, C. C., Buyukkececi, Z., & Çineli, B. (2025). Mapping modern kinship networks: First results from the KINMATRIX survey. *Journal of Marriage and Family*, 87(2), 478–504. <https://doi.org/10.1111/jomf.13049>
- Macapagal, K., Coventry, R., Arbeit, M. R., Fisher, C. B., & Mustanski, B. (2017). “I Won’t Out Myself Just to Do a Survey”: Sexual and Gender Minority Adolescents’ Perspectives on the Risks and Benefits of Sex Research. *Archives of Sexual Behavior*, 46(5), 1393–1409. <https://doi.org/10.1007/s10508-016-0784-5>
- Martin, J. I., & D’Augelli, A. R. (2007). Timed Lives: Cohort and Period Effects in Research on Sexual Orientation and Gender Identity. In *Handbook of Research with Gay, Lesbian, Bisexual & Transgender Populations* (pp. 190–207). Taylor & Francis Ltd. <https://research.ebsco.com/linkprocessor/plink?id=609a2306-69bf-3e7c-b18e-b31f88fbb50a>
- Matthews, D. D., Blosnich, J. R., Farmer, G. W., & Adams, B. J. (2014). Operational Definitions of Sexual Orientation and Estimates of Adolescent Health Risk Behaviors. *LGBT Health*, 1(1), 42–49. <https://doi.org/10.1089/lgbt.2013.0002>
- Max Planck Institute for Demographic Research, University of California, Berkeley, & French Institute for Demographic Studies. (2026). *Human Mortality Database* [Dataset]. www.mortality.org
- Max Planck Institute for Demographic Research & Vienna Institute of Demography. (2026). *Human Fertility Database* [Dataset]. www.humanfertility.org
- McPherson, M., Smith-Lovin, L., & Cook, J. M. (2001). Birds of a Feather: Homophily in Social Networks. *Annual Review of Sociology*, 27, 415–444.
- Mittleman, J. (2023). Sexual Fluidity: Implications for Population Research. *Demography*, 60(4), 1257–1282. <https://doi.org/10.1215/00703370-10898916>

- Moore, M. R., & Stambolis-Ruhstorfer, M. (2013). LGBT Sexuality and Families at the Start of the Twenty-First Century. *Annual Review of Sociology*, 39(Volume 39, 2013), 491–507. <https://doi.org/10.1146/annurev-soc-071312-145643>
- Neidorf, S., & Morin, R. (2007). *Four-in-Ten Americans Have Close Friends or Relatives Who are Gay*. Pew Research Center. <https://www.pewresearch.org/social-trends/2007/05/22/fourinten-americans-have-close-friends-or-relatives-who-are-gay/>
- Park, N. K., Kazyak, E., & Slauson-Blevins, K. (2016). How Law Shapes Experiences of Parenthood for Same-Sex Couples. *Journal of GLBT Family Studies*, 12(2), 115–137. <https://doi.org/10.1080/1550428X.2015.1011818>
- Pillard, R. C. (1986). Evidence of Familial Nature of Male Homosexuality. *Archives of General Psychiatry*, 43(8), 808. <https://doi.org/10.1001/archpsyc.1986.01800080094012>
- Reczek, R. (2016). Ambivalence in Gay and Lesbian Family Relationships. *Journal of Marriage and Family*, 78(3), 644–659. <https://doi.org/10.1111/jomf.12308>
- Reczek, R. (2025). Mapping the Cultural Repertoires of Family Estrangement: A New Theory of Democratized Kinship. *Social Problems*, spaf026. <https://doi.org/10.1093/socpro/spaf026>
- Reczek, R., Stacey, L., & Dunston, C. (2022). Friend, foe, or forget ‘em?: The quality of LGBTQ adult sibling relationships. *Journal of Marriage and Family*, 84(2), 415–437. <https://doi.org/10.1111/jomf.12821>
- Reczek, R., Stacey, L., & Thomeer, M. B. (2023). Parent–adult child estrangement in the United States by gender, race/ethnicity, and sexuality. *Journal of Marriage and Family*, 85(2), 494–517. <https://doi.org/10.1111/jomf.12898>
- Robinson, B. A., Stone, A. L., & Webb, J. M. (2023). Aunties, Aunts, and Tías: The Forgotten Othermother Supporting and Housing LGBTQ Youth. *Socius: Sociological Research for a Dynamic World*, 9, 23780231231217827. <https://doi.org/10.1177/23780231231217827>
- Rosenfeld, M. J. (2017). Moving a Mountain: The Extraordinary Trajectory of Same-Sex Marriage Approval in the United States. *Socius: Sociological Research for a Dynamic World*, 3, 2378023117727658. <https://doi.org/10.1177/2378023117727658>
- Ruggles, S. (1987). *Prolonged connections: The rise of the extended family in nineteenth-century England and America* (1. print). Univ. of Wisconsin Press.
- Saperstein, A. (2025). Recognizing Identity Fluidity in Demographic Research. *Population and Development Review*, 51(1), 519–538. <https://doi.org/10.1111/padr.12670>
- Scherrer, K. S. (2016). Gay, Lesbian, Bisexual, and Queer Grandchildren’s Disclosure Process With Grandparents. *Journal of Family Issues*, 37(6), 739–764. <https://doi.org/10.1177/0192513X14526874>
- Scherrer, K. S., Kazyak, E., & Schmitz, R. (2015). Getting “Bi” in the Family: Bisexual People’s Disclosure Experiences. *Journal of Marriage and Family*, 77(3), 680–696. <https://doi.org/10.1111/jomf.12190>
- Silva, T. (2019). Straight Identity and Same-Sex Desire: Conservatism, Homophobia, and Straight Culture. *Social Forces*, 97(3), 1067–1094. <https://doi.org/10.1093/sf/soy064>

- Silva, T. (2025). Sexual Identities: The State of the Field and Future Directions. *Sex & Sexualities*, 1(1), 172–181. <https://doi.org/10.1177/3033371251331903>
- Silva, T., & Evans, C. R. (2022). How Do Adolescent Social Determinants and Social Contexts Shape Adult Sexual Identification? *Social Problems*, 69(2), 817–840. <https://doi.org/10.1093/socpro/spaa074>
- Song, X., & Caswell, H. (2022). The Role of Kinship in Racial Differences in Exposure to Unemployment. *Demography*, 59(4), 1325–1352. <https://doi.org/10.1215/00703370-10057831>
- Stone, A. L., Pride, C., Adams, B., & Salcido, R. (2022). Bisexual Aunts, Gay Cousins: How LGBTQ Family Members Help LGBTQ Youth Navigate Family Heteronormativity. *Socius: Sociological Research for a Dynamic World*, 8, 23780231221117145. <https://doi.org/10.1177/23780231221117145>
- Van Damme, M., Alburez-Gutierrez, D., & Castro Torres, A. F. (2025). Research Note: Estimating Kinship Size of Older Adults in Europe With Models and Surveys. *Demography*, 11961236. <https://doi.org/10.1215/00703370-11961236>
- Verdery, A. M. (2015). Links Between Demographic and Kinship Transitions. *Population and Development Review*, 41(3), 465–484. JSTOR. <https://doi.org/10.1111/j.1728-4457.2015.00068.x>
- Verdery, A. M., & Margolis, R. (2017). Projections of white and black older adults without living kin in the United States, 2015 to 2060. *Proceedings of the National Academy of Sciences*, 114(42), 11109–11114. <https://doi.org/10.1073/pnas.1710341114>
- Weston, K. (1997). *Families we choose: Lesbians, gays, kinship*. Columbia University Press.
- Williams, I., & Alburez-Gutierrez, D. (2023). *DemoKin: Estimate Population Kin Distribution* (p. 1.0.3) [Dataset]. CRAN. <https://doi.org/10.32614/CRAN.package.DemoKin>

TABLES AND FIGURES

Table 1 Share of population with one or more queer kin, by kin sexuality, kin genealogical closeness, and Focal age

Age of Focal and Kin Sexuality	Any kin ^a (%, 95% CI)	Genealogical closeness	
		Nuclear kin ^b (%, 95% CI)	Extended kin ^c (%, 95% CI)
Focal aged 15 (Born in 2009)			
Queer ^d	74.8 (71.6, 77.7)	17.9 (16.4, 19.4)	69.3 (66, 72.3)
<i>Bisexual</i>	53.8 (49.8, 57.6)	11.4 (10.2, 12.6)	47.9 (44.1, 51.5)
<i>Gay/Lesbian</i>	28.3 (24.4, 32)	3.6 (3, 4.3)	25.6 (22.1, 28.9)
<i>Other</i>	17.5 (13.7, 21.3)	2.5 (1.9, 3.1)	15.4 (12, 18.7)
Age 30 (1994)			
Queer	87.4 (85, 89.5)	29.1 (27.1, 31.1)	82.3 (79.5, 84.7)
<i>Bisexual</i>	67 (62.8, 70.7)	17.9 (16.2, 19.5)	59.8 (55.6, 63.6)
<i>Gay/Lesbian</i>	42.8 (37.9, 47.4)	7.7 (6.6, 8.8)	38 (33.5, 42.3)
<i>Other</i>	26 (20.6, 31.1)	4.6 (3.6, 5.6)	22.5 (17.7, 27)
Age 45 (1979)			
Queer	88.2 (85.7, 90.2)	36.5 (33.9, 39.1)	81.4 (78.4, 84)
<i>Bisexual</i>	66.1 (61.6, 70.1)	22.2 (20, 24.4)	56.4 (52, 60.4)
<i>Gay/Lesbian</i>	46.5 (41.3, 51.3)	10.7 (9, 12.3)	40.1 (35.5, 44.5)
<i>Other</i>	27.9 (21.8, 33.5)	6.3 (4.8, 7.8)	23.1 (17.9, 27.9)
Age 60 (1964)			
Queer	83.8 (80.9, 86.3)	35.5 (33.1, 37.9)	74.9 (71.5, 77.9)
<i>Bisexual</i>	57.9 (53.3, 62)	18.9 (17, 20.8)	48.1 (43.7, 52.1)
<i>Gay/Lesbian</i>	44.6 (39.7, 49.2)	13.4 (11.7, 15)	36.1 (31.8, 40.2)
<i>Other</i>	25.2 (19.5, 30.6)	6.6 (5.1, 8.1)	19.9 (15.2, 24.5)
Age 75 (1949)			
Queer	77.6 (74, 80.7)	16.9 (15.2, 18.7)	73 (69.4, 76.2)
<i>Bisexual</i>	50.6 (45.9, 55)	6.1 (5, 7.1)	47.4 (43, 51.5)
<i>Gay/Lesbian</i>	38.3 (33.5, 42.7)	7.8 (6.7, 9)	33 (28.8, 37.1)
<i>Other</i>	21.7 (16.3, 26.9)	3.7 (2.6, 4.8)	18.7 (14, 23.2)

Notes: Estimates based on 2024 kinship structure projected by time-variant, two-sex kinship matrix model and pooled 2021-2024 rates of queer identification. Kin are inclusive of the 18-87 age range. "CI" = 95% confidence interval calculated from the survey estimates only as the kinship model does not account for uncertainty in the number of kin. Categories are not mutually exclusive; estimates are irrespective of if, e.g., Focal has queer extended kin (for queer nuclear cell), or if Focal has bisexual nuclear kin (for gay/lesbian extended kin cell). ^a Nuclear V extended kin. ^b Siblings V children, exclusive of parents and partners. ^c aunts/uncles V cousins V nieces/nephews V grandchildren. ^d Bisexual V gay/lesbian V other non-straight.

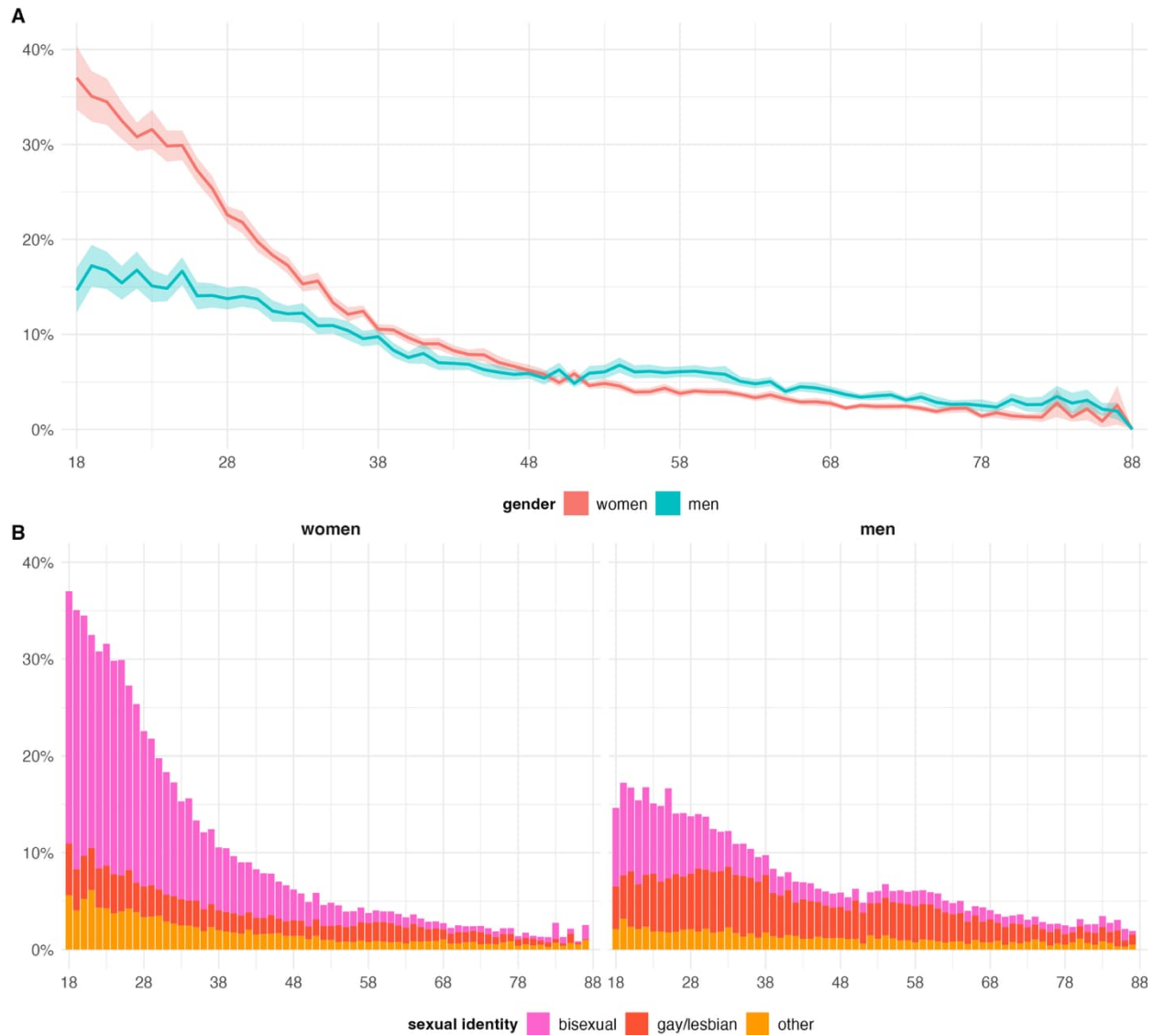


Figure 1 Population share with queer identity in the United States, by gender and age

Notes: “Queer” refers to lesbian, gay, bisexual, and other non-straight identities. Panel A presents author’s calculations of the age-sex specific share identifying as queer using survey-weighted data from Phase 3.2–Phase 4.2 of the Household Pulse Survey for ages 18-87. Panel B disaggregates this share into the share identifying as bisexual, gay/lesbian, or other non-straight identity. Survey weights are applied, 95% confidence intervals shaded.

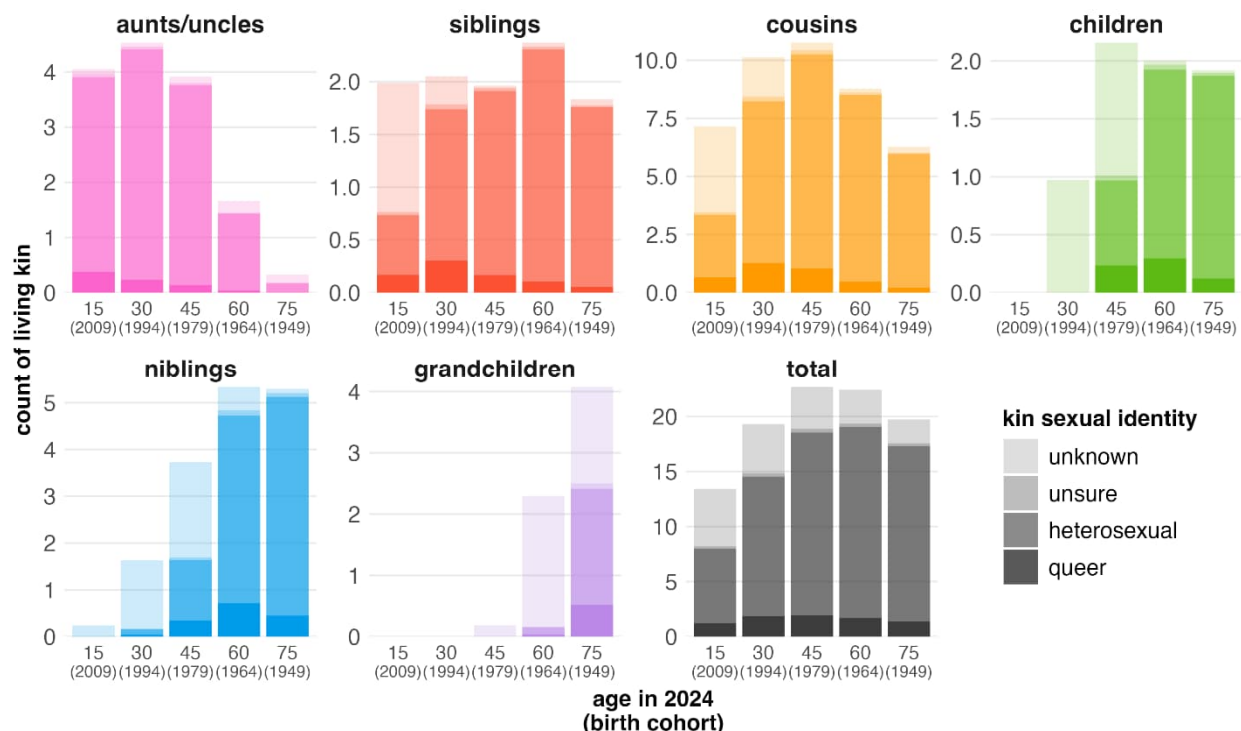


Figure 2 Composition of kinship network by kin sexual identity and Focal age

Notes: Estimates for an average person aged 15, 30, 45, 60, and 75 based on projected 2024 kinship structure and pooled 2021-2024 rates of queer identification. The corresponding birth cohort of Focal is written in parentheses beneath their age in 2024. Within each bar, the darker shade indicates the expected number of queer kin, the second darkest shade indicates the expected number of heterosexual (straight) kin, the second lightest shade captures kin who are unsure about their identity, and the lightest shade captures “unknown” kin for whom we cannot estimate their identity due to being out of the age range for the survey or due to survey nonresponse. Queer includes lesbian/gay, bisexual, and other non-straight identities. “Nibbling” = niece or nephew. *Source:* Estimates based on two-sex, time-variant kinship matrix model. Demographic data are drawn from the Human Mortality Database and Human Fertility Database. Survey data on queer identification are drawn from Phase 3.2–Phase 4.2 of the Household Pulse Survey for ages 18-87. Survey weights are applied to ensure the findings are nationally representative.

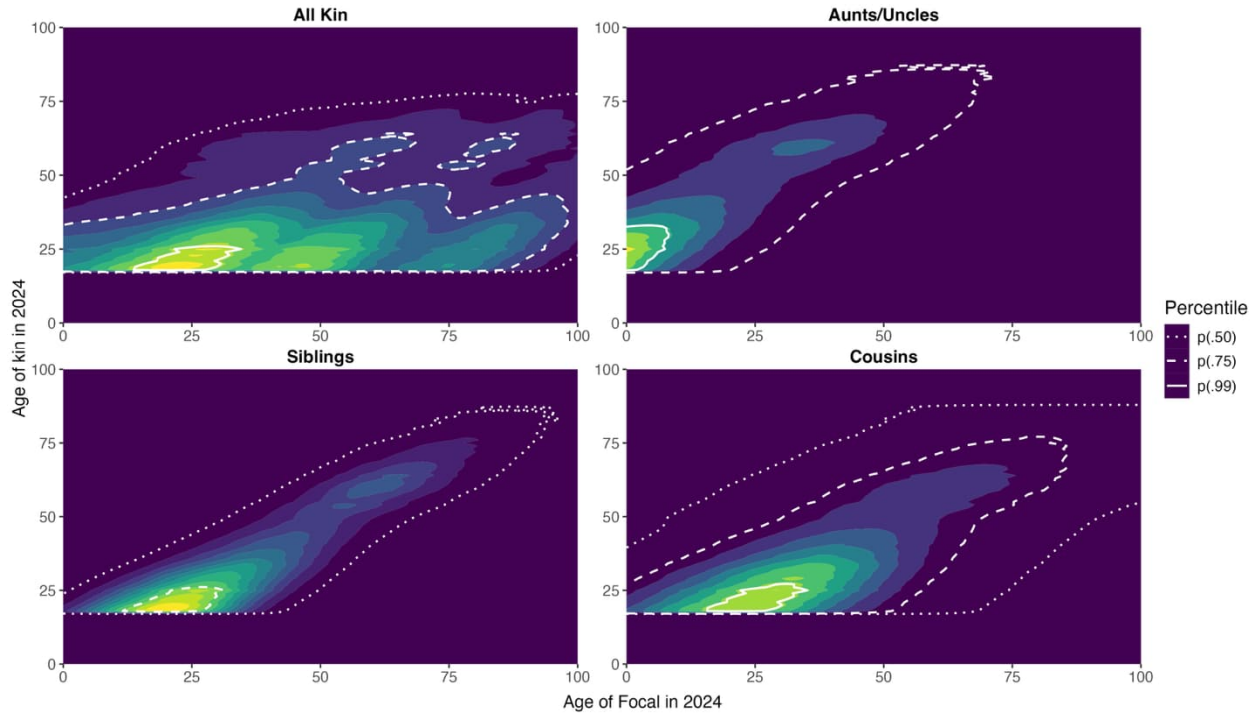


Figure 3 Joint age-specific distribution of queer kin by age of Focal, age of kin, and kin type

Notes: Estimates based on projected 2024 kinship structure and pooled 2021-2024 rates of queer identification. Queer includes lesbian/gay, bisexual, and other non-straight identities. Lighter shades indicate relatively more queer kin, whereas darker shades indicate relatively fewer. Queer peers are individuals within a five-year age range of Focal, or 10 or more years older for aunts/uncles. *Source:* Estimates based on two-sex, time-variant kinship matrix model. Demographic data are drawn from the Human Mortality Database and Human Fertility Database. Survey data on queer identification are drawn from Phase 3.2–Phase 4.2 of the Household Pulse Survey for ages 18-87. Survey weights are applied to ensure the findings are nationally representative.

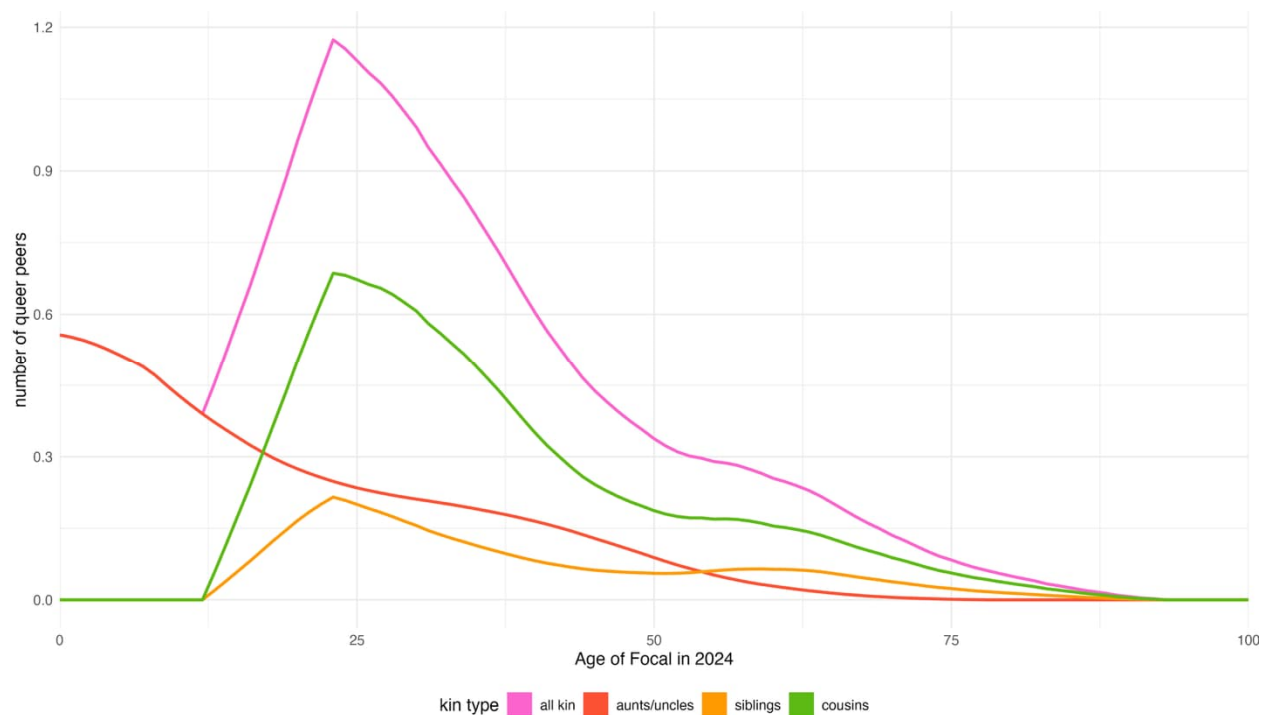


Figure 4 Distribution of queer peers, by age of Focal and kin type

Notes: Estimates based on projected 2024 kinship structure and pooled 2021-2024 rates of queer identification. Queer includes lesbian/gay, bisexual, and other non-straight identities. Queer peers are individuals within a five-year age range of Focal, or 10 or more years older for aunts/uncles. *Source:* Estimates based on two-sex, time-variant kinship matrix model. Demographic data are drawn from the Human Mortality Database and Human Fertility Database. Survey data on queer identification are drawn from Phase 3.2–Phase 4.2 of the Household Pulse Survey for ages 18-87. Survey weights are applied to ensure the findings are nationally representative.

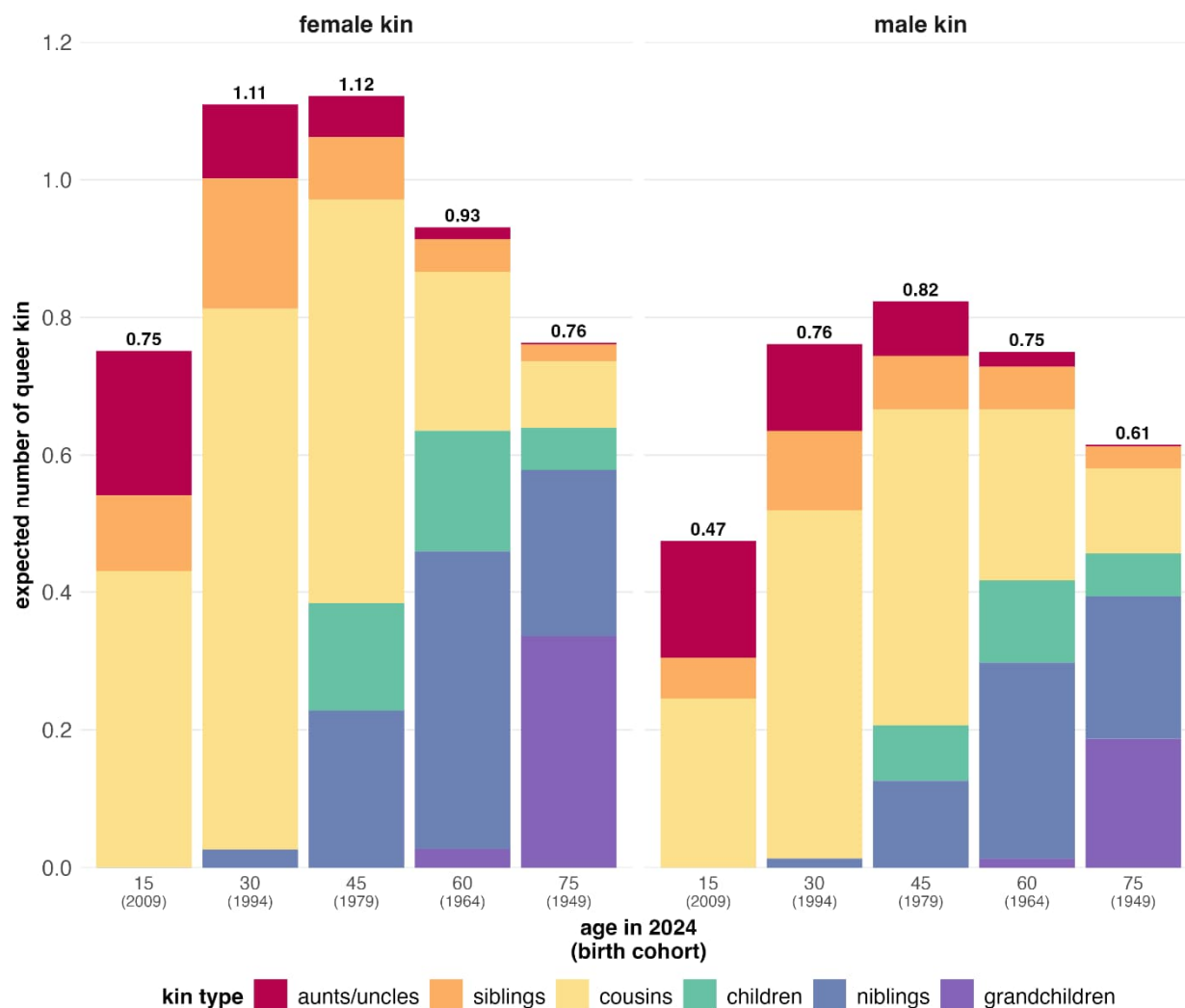


Figure 5 Expected number of queer female and male kin, by kin type and Focal age

Notes: Estimates for an average person aged 15, 30, 45, 60, and 75 based on projected 2024 kinship structure and pooled 2021-2024 rates of queer identification. The corresponding birth cohort of Focal is written in parentheses beneath their age in 2024. Bolded numbers at the top of each stacked bar represent the sum of queer kin at each age of Focal. Queer includes lesbian/gay, bisexual, and other non-straight identities. “Nibling” = niece or nephew. Survey data on queer identification are drawn from Phase 3.2–Phase 4.2 of the Household Pulse Survey. Survey weights are applied to ensure the findings are nationally representative.

Queer family matters: Estimates of queer kinship in the United States

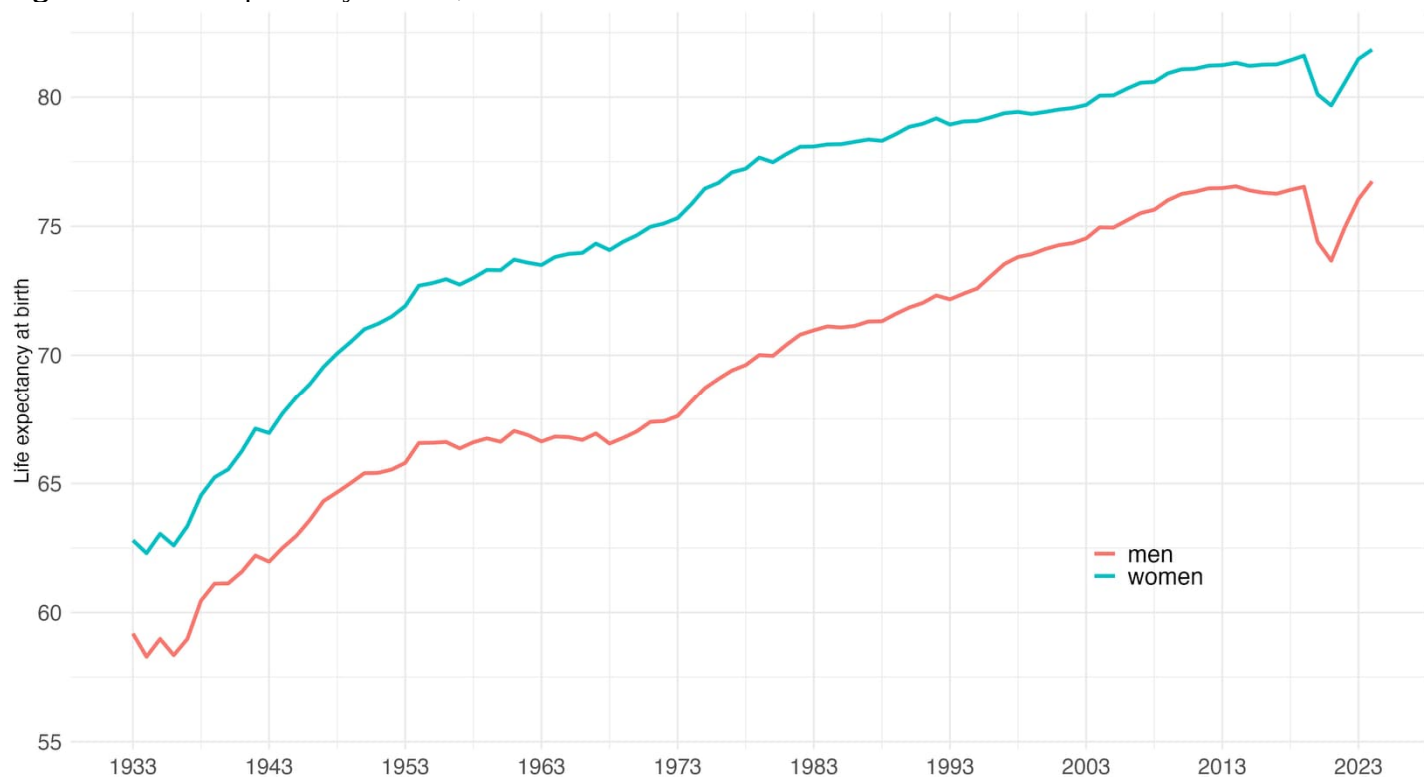
Online Appendix

A. Demographic inputs

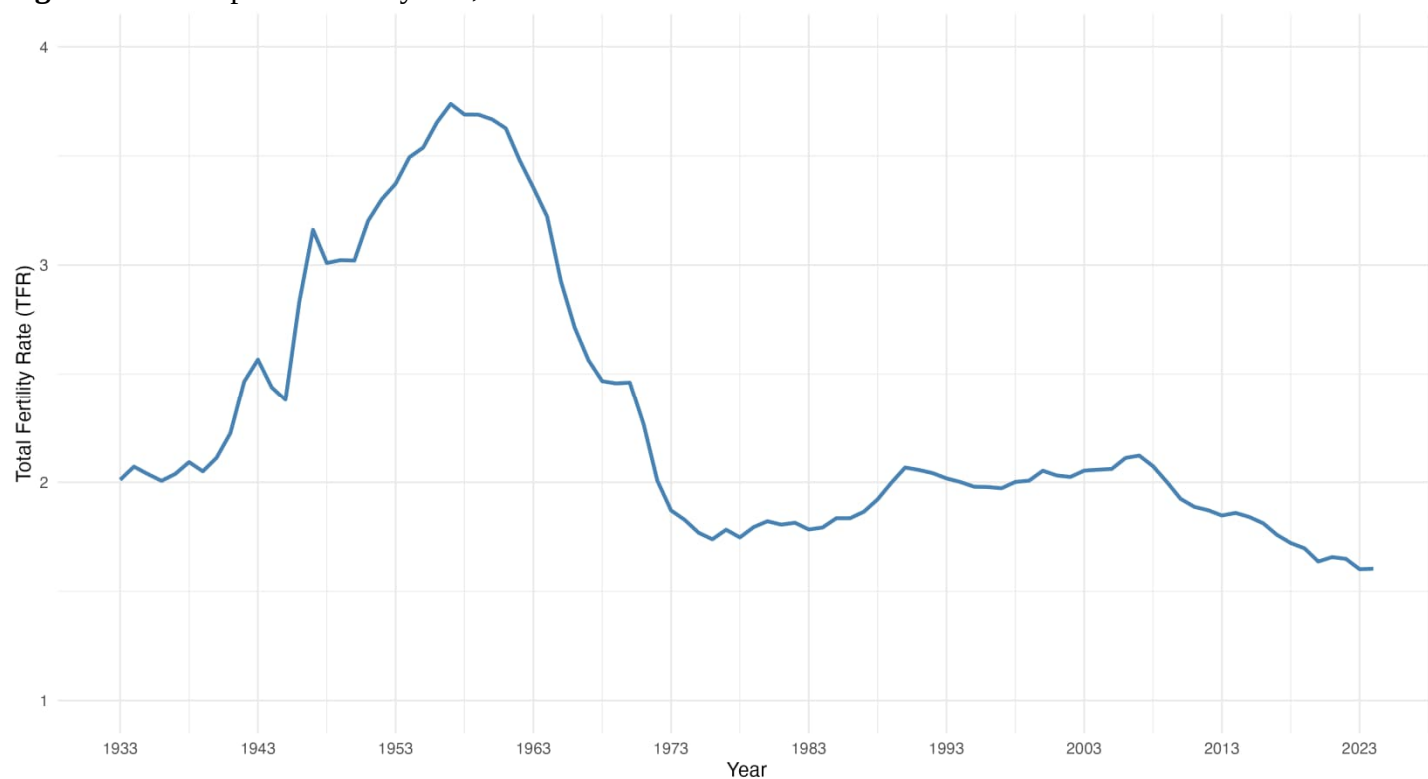
The 1933 rates are applied to all prior years (i.e., a stable population is assumed), which should not affect our results as the fertility transition in the United States began over 100 years prior (Hacker & Roberts, 2020) and should only affect ancestors (e.g., parents) (Alburez-Gutierrez et al., 2023), who are not measured. Figures S1-S4 show the demographic inputs to the kinship matrix model.

Although the Human Fertility Database provides age-specific fertility rates for females between 1933-2024, the Human Fertility Collection only has age-specific fertility rates for males between 1969-2015. We impute the missing male fertility rates (1933-1968 and 2016-2024) using the regression-based approach developed by Abrams et al. (2025). However, for missing values of the dynamic age difference term, we impute a value of three as opposed to two years (Schoumaker, 2019). The results of this imputation for the year 2024 can be seen in Figure S4. This approach yielded substantively similar findings to if the whole male fertility schedule had been adjusted rightward three years.

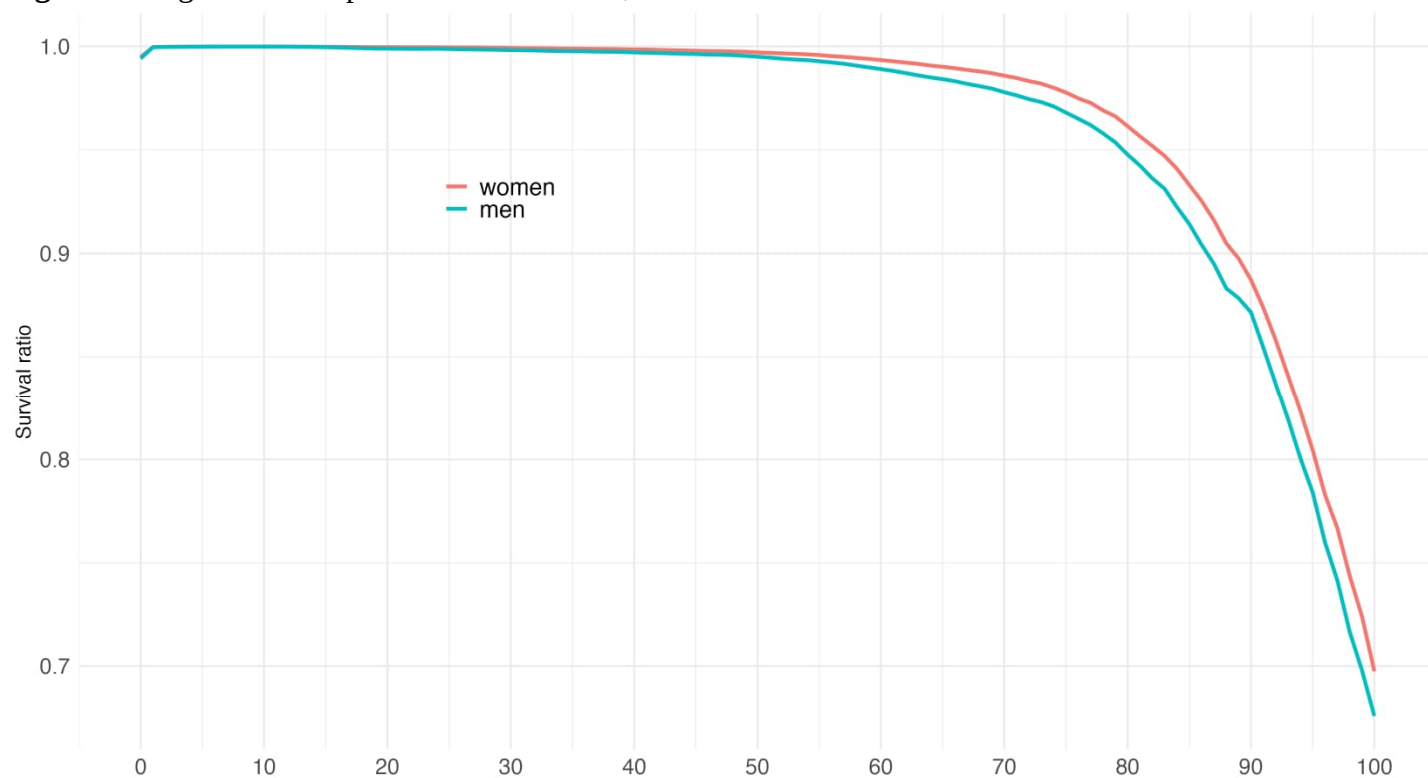
Figure S1. Life expectancy at birth, 1933-2024



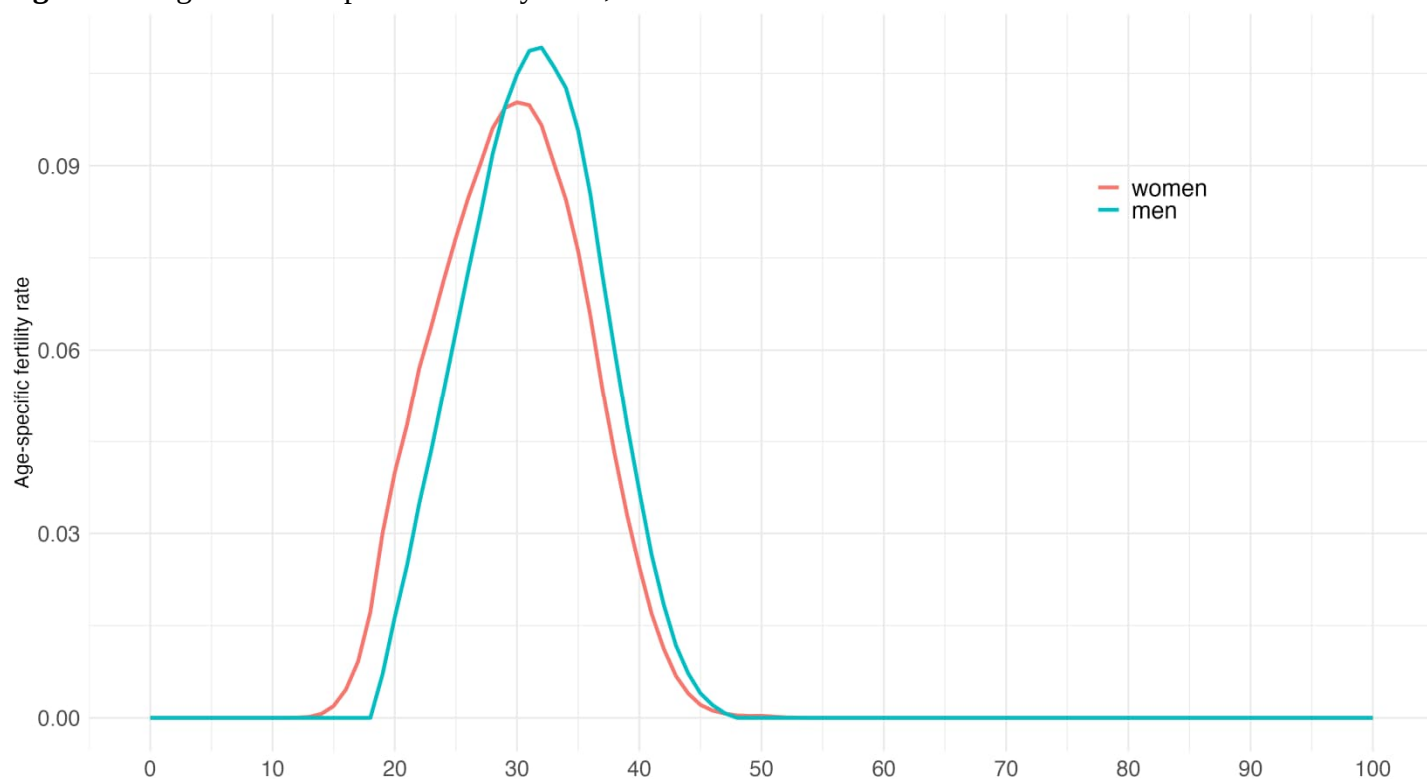
Note: Figure shown to explain demographic trends over time. Data from the Human Mortality Database.

Figure S2. Total period fertility rate, 1933-2024

Note: Figure shown to explain demographic trends over time. Authors' calculation based on the Human Fertility Database.

Figure S3. Age- and sex-specific survival ratios, 2024

Note: Example of period values used in analysis. Authors' calculation based on Human Mortality Database.

Figure S4. Age- and sex-specific fertility rates, 2024

Note: Example of period values used in analysis. Authors' calculation based on the Human Fertility Database.

B. Comparability of queer sexual identity measure across surveys

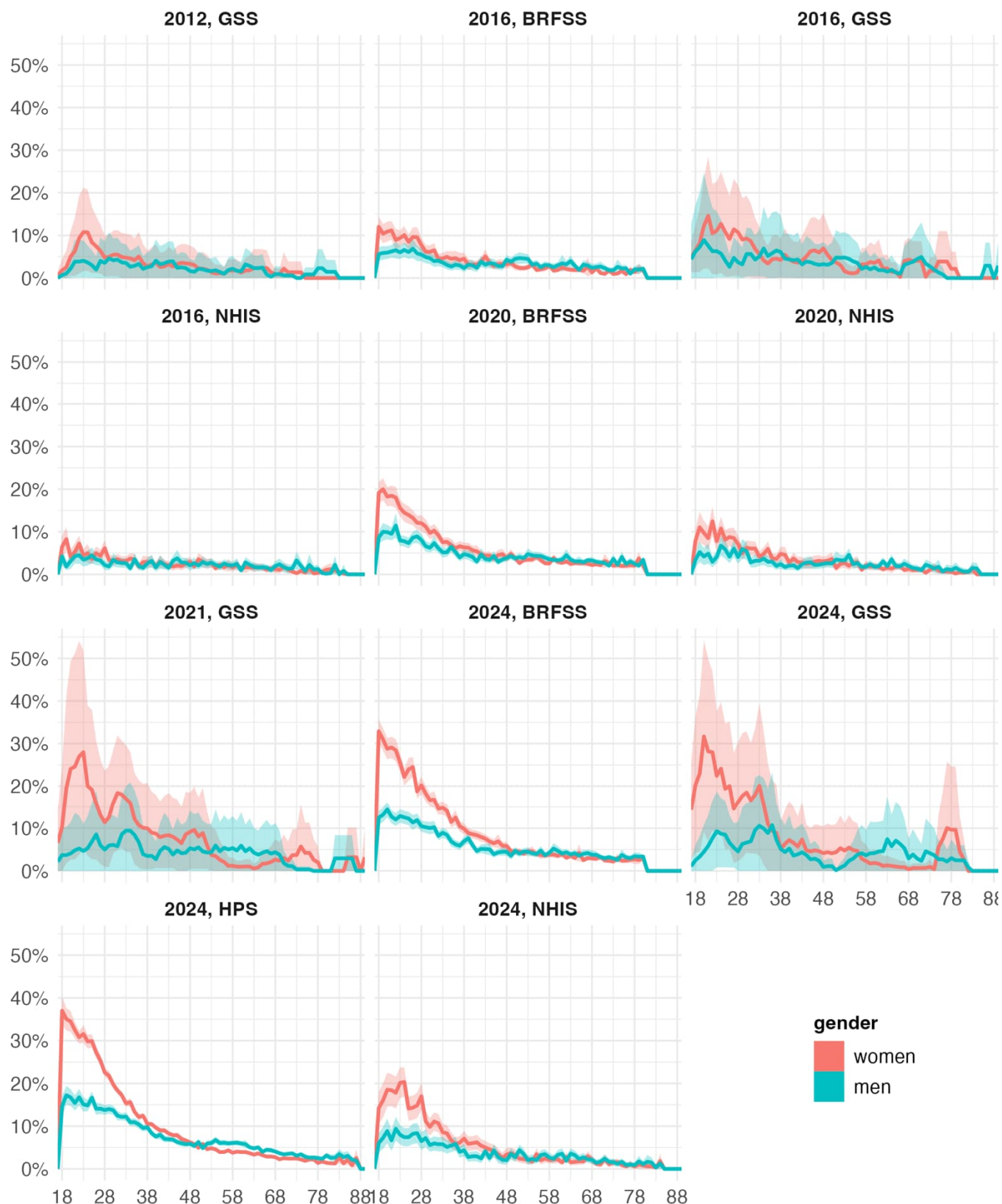
Table S1 compares the Household Pulse Survey (HPS) measures to those in the National Health Interview Survey (NHIS), General Social Survey (GSS), and Behavioral Risk Factor Surveillance System (BRFSS).

Table S1. Queer measure construction across survey waves

	Queer	Straight	Unsure	Unknown
HPS Phase 3.2-3.10	Gay or lesbian Bisexual Something else	Straight, that is not gay or lesbian	I don't know	Non-response
HPS 4.0-4.2	Gay or lesbian Bisexual I use a different term	Straight, that is not gay or lesbian	-	Non-response
NHIS 2013-2018	Lesbian or gay Bisexual Something else	Straight, that is, not lesbian or gay	I don't know the answer	Unknown-refused Unknown-not ascertained
NHIS 2019-2023	Gay/lesbian Bisexual Something else	Straight, that is not gay/lesbian	I do not know the answer Do not know	Refused
GSS 2008-2024	Gay, lesbian, or homosexual Bisexual	Heterosexual or straight	Do not Know/Cannot Choose	Non-response Skipped on web
BRFSS 2014-2017	Lesbian or gay Bisexual Other	Straight	Don't know/Not Sure	Refused Non-response
BRFSS 2018-2024	Gay Bisexual Something else	Straight, that is, not gay	I don't know the answer	Refused Non-response

Figure S5 compares the age-sex schedule of these four data sources across waves. This figure also shows how the share of queer sexual identity has increased over time, driven by increases by young people, particularly women.

BRFSS produces a similar queer schedule to what is estimated by the HPS. However, the HPS is not limited to select states like the BRFSS—a problem for identifying individuals with both a queer sexual *and* gender identity (Julian et al., 2024)—and estimates from the HPS in BRFSS-fielded states are lower than the national average (Deng & Watson, 2023), although states that do not field this module tend to have smaller queer populations (Brady et al., 2025). The age range is also more limited. Given the concordance of the BRFSS estimates with the HPS, the queer kin estimates are only moderately smaller due to the smaller age range and slightly lower estimates among younger people in BRFSS. The GSS, by contrast is a poor comparison to the HPS: it has large confidence intervals and lacks a smooth pattern, even after applying a 5-year smoothing. This notwithstanding, the GSS produces age-sex schedules similar to the BRFSS. The NHIS produces a queer schedule that is lower than what is estimated by the HPS or BRFSS, particularly among younger ages, where the estimate is about half of the HPS. This underestimate is elsewhere noted (Deng & Watson, 2023; Julian et al., 2024), and may be due to the face-to-face interview format which could discourage sexual identity disclosure due to social desirability bias or fear of discrimination (Lee et al., 2018). After age 48 however, the estimates are quite similar. Several other considerations guide our choice to forgo the NHIS: larger standard errors, second, smaller age range, and dissimilar age-sex patterns.

Figure S5. Population share with queer identity in the United States, by gender and age

Note: Authors' calculations using survey-weighted data from the 2014-2016, 2017-2020, and 2021-2024 waves of the Behavioral Risk Surveillance System (ages 18-80); 2008-2012, 2013-2016, 2017-2021, and 2022-2024 waves of the General Social Survey (ages 18-84); Phase 3.2-4.2 of the Household Pulse Survey (ages 18-87); 2013-2016, 2017-2020, and 2021-2024 waves of the National Health Interview Survey (ages 18-85). Survey weights are applied, 95% confidence intervals shaded. "Queer" refers to lesbian, gay, bisexual, and other non-straight identities.

To further give confidence to our estimates, particularly with respect to change across time, we benchmark our findings from the HPS to those from the BRFSS, GSS, and NHIS. Because these four datasets all have different maximum ages, we restrict the comparison to kin aged 18-80. Figure S6 demonstrates that the HPS estimates of queer kin for a Focal aged 30 in 2024 are quite similar to those estimated using the BRFSS: around 12.5% of all kin from the HPS compared to 10.3% from the BRFSS, a difference of about two percentage points that is also seen in the estimates for female and male kin. Both the HPS and BRFSS estimate a higher share relative to the GSS or NHIS, likely because they underestimate queer identity among young people (Julian et al., 2024).

This figure also demonstrates that the share of an individual's kinship network who identifies as queer has been increasing steadily over the past few years, rising from around 3.5% in 2012 using the GSS estimates to between 6.6% and 12.5% using the NHIS and HPS, respectively. Additionally, queer female kin constitute a larger share of the kinship network than queer male kin, and the difference has grown across time.

Additional analyses (not shown) replicate Figure S6 for lesbian and gay kin only. Here, the estimates are extremely concordant for female kin, suggesting that the measurement of lesbians is quite similar across sources, but the HPS estimates a greater number of gay male kin, possibly because gay men feel more confident in disclosing their identity in an online survey versus face-to-face or over the phone (Deng & Watson, 2023). This finding also suggests that the higher share seen for female kin below may be due to more women reporting their identity as bisexual or something else in the HPS compared to the three other sources, again potentially due to survey mode.

Figure S6. Estimates of queer kin as share of total kin across four data sources



C. Supplemental Findings

Youth Risk Behavior Survey (YRBS)

We attempt to expand the age range of our queer schedule by using the 2021 and 2023 waves of the Youth Risk Behavior Survey (YRBS), for which we are able to use data from respondents aged 14-17. We use data from the 2021 and 2023 waves of the YRBS, which has been fielded since 1991 to high school students regardless of public or private status. In addition to student demographics, a sexual identity question has been included since 2015 with an update to response options in 2021. We exclude respondents aged 12 and 13 due to small sample size, and do not use estimates from 18-year-olds in favor of those from the HPS. Starting in 2021, the sexual identity measurement was as follows: *Which of the following best describes you? Heterosexual (straight), gay or lesbian, bisexual, I describe my sexual identity some other way, I am not sure about my sexual identity (questioning), I do not know what this question is asking.* These findings contribute to greater research on queer adolescents who represent an understudied group in the literature (Macapagal et al., 2017).

Table S2. Share of population with one or more queer kin, by kin sexuality, kin genealogical closeness, and Focal age

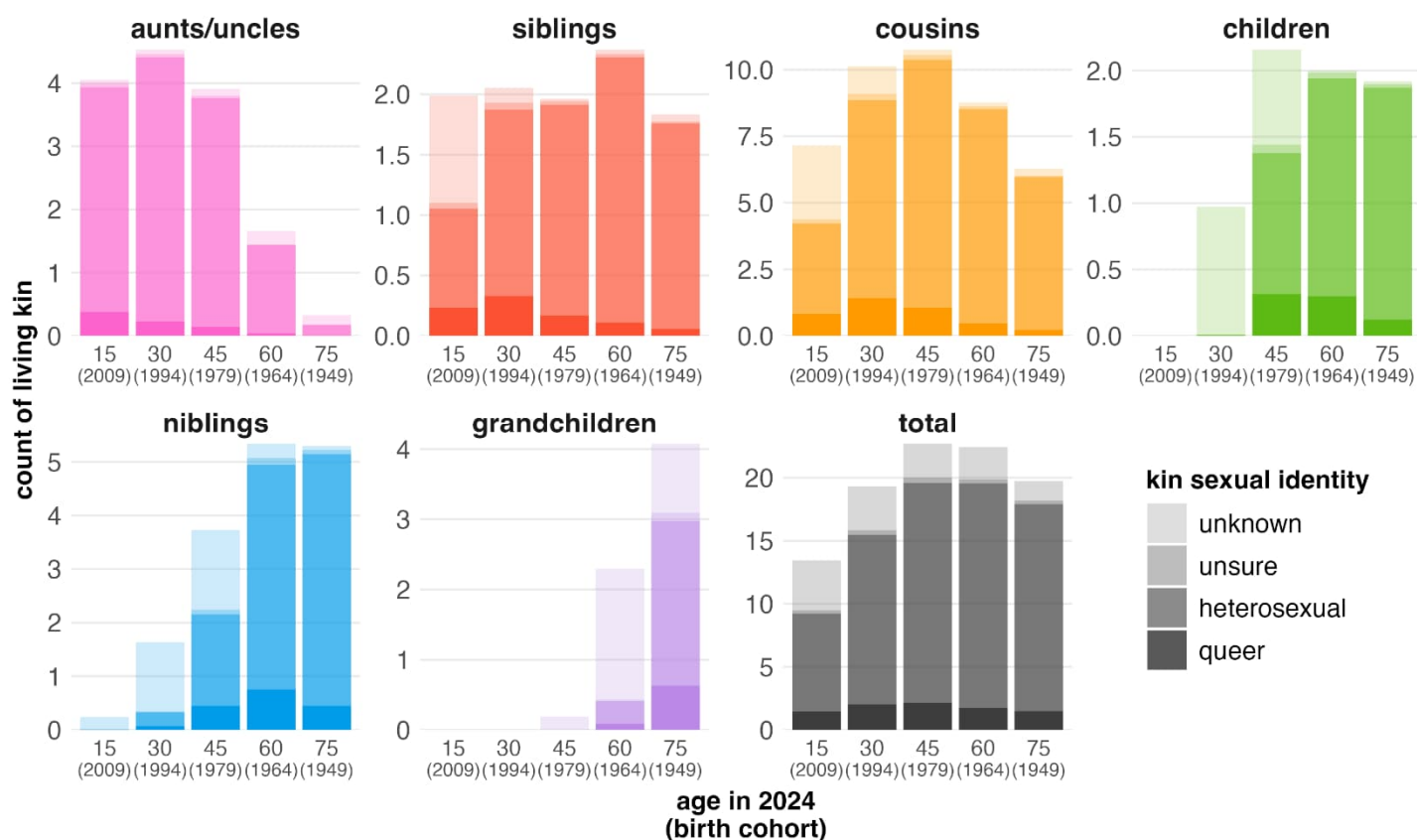
Age of Focal and Kin Sexuality	Any kin ^a (%, 95% CI)	Genealogical closeness	
		Nuclear kin ^b (%, 95% CI)	Extended kin ^c (%, 95% CI)
Focal aged 15 (Born in 2008)			
Queer ^d	81 (77.8, 83.8)	23.3 (21.2, 25.4)	75.2 (71.8, 78.2)
Bisexual	61.2 (56.7, 65.3)	15.1 (13.4, 16.8)	54.3 (50, 58.3)
Gay/Lesbian	31.4 (26.8, 35.8)	4.6 (3.7, 5.6)	28.1 (24, 32)
Other	21.3 (16.5, 25.8)	3.6 (2.7, 4.5)	18.3 (14.2, 22.3)
Focal aged 30 (Born in 1993)			
Queer	90.1 (87.8, 91.9)	31.1 (28.9, 33.3)	85.6 (82.8, 87.9)
Bisexual	71.4 (67.1, 75.2)	19.3 (17.5, 21.2)	64.5 (60.1, 68.5)
Gay/Lesbian	45 (39.6, 49.9)	8.1 (6.9, 9.4)	40.1 (35.2, 44.7)
Other	28.8 (22.8, 34.3)	5 (3.9, 6.2)	25 (19.6, 30)
Focal aged 45 (Born in 1978)			
Queer	90.7 (88.4, 92.6)	42.4 (39.2, 45.4)	83.9 (81, 86.4)
Bisexual	70.9 (66.3, 74.9)	26.7 (23.9, 29.4)	60.3 (55.7, 64.4)
Gay/Lesbian	48.5 (42.8, 53.6)	12 (10, 14)	41.4 (36.5, 46)
Other	30.7 (24, 36.8)	7.8 (5.9, 9.7)	24.9 (19.3, 30.1)
Focal aged 60 (Born in 1963)			
Queer	85.6 (82.8, 88)	35.6 (33.1, 37.9)	77.7 (74.3, 80.7)
Bisexual	60.9 (56.1, 65.1)	18.9 (17, 20.8)	51.7 (47.2, 56)
Gay/Lesbian	45.7 (40.6, 50.4)	13.4 (11.7, 15.1)	37.3 (32.7, 41.6)
Other	26.6 (20.6, 32.2)	6.6 (5.1, 8.1)	21.4 (16.3, 26.3)
Focal aged 75 (Born in 1948)			
Queer	80.5 (77, 83.5)	16.9 (15.2, 18.7)	76.5 (72.9, 79.7)
Bisexual	54.7 (49.7, 59.2)	6.1 (5, 7.1)	51.8 (47.1, 56.1)
Gay/Lesbian	39.6 (34.5, 44.2)	7.9 (6.7, 9)	34.4 (29.8, 38.7)
Other	23.4 (17.6, 28.9)	3.7 (2.6, 4.8)	20.5 (15.4, 25.3)

Notes: Kin are inclusive of the 18-87 age range. "CI" = 95% confidence interval. Categories are not mutually exclusive; estimates are irrespective of if, e.g., Focal has queer extended kin (for queer nuclear cell), or if Focal has bisexual nuclear kin (for gay/lesbian extended kin cell). ^a Nuclear V extended kin. ^b Siblings V children, exclusive of parents and partners. ^c aunts/uncles V cousins V nieces/nephews V grandchildren. ^d Bisexual V gay/lesbian V other non-straight.

Figure S7. Population share with queer identity in the United States, by gender and age (YRBS and HPS)

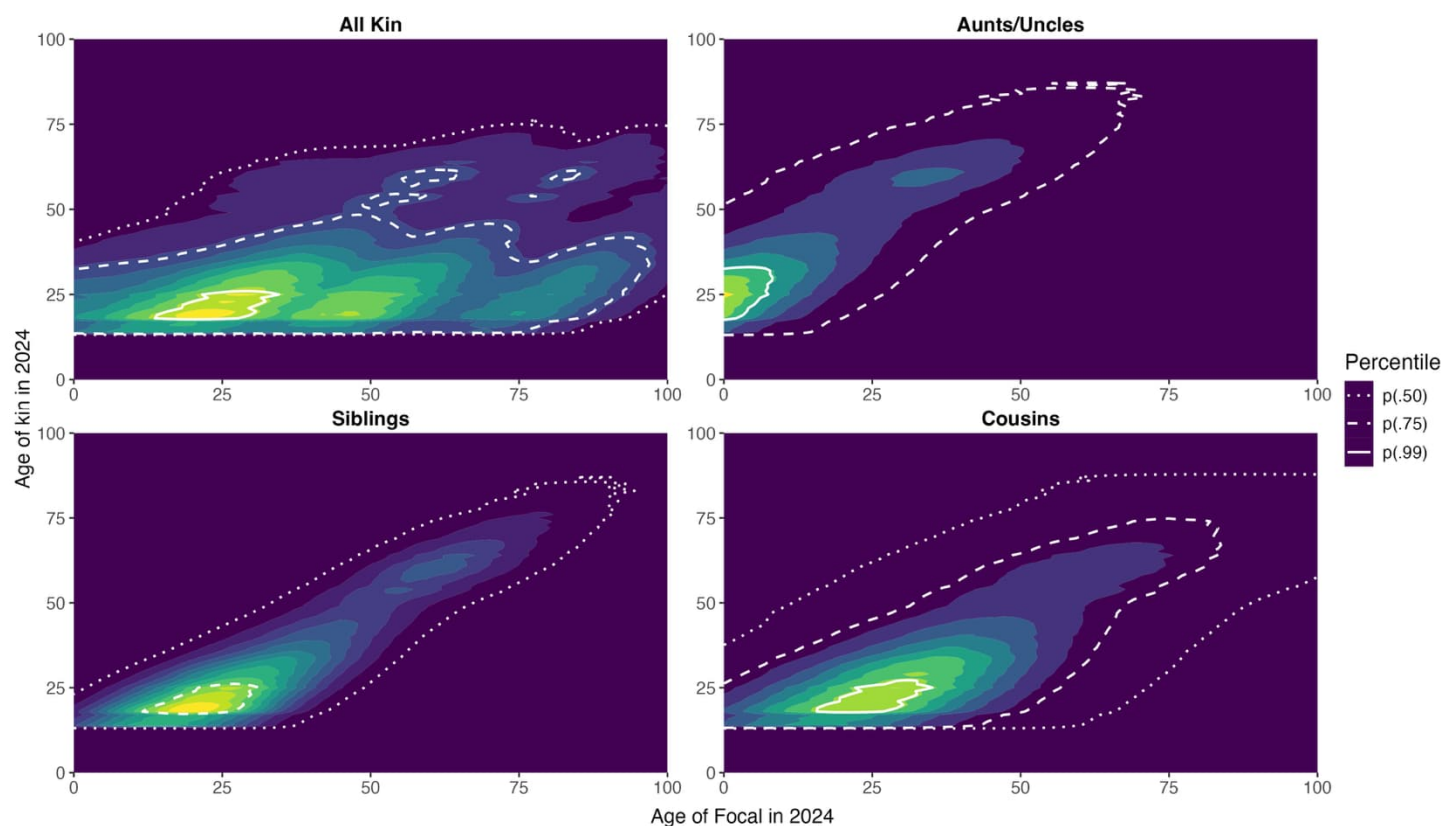


Note: Authors' calculations using survey-weighted data from the 2021 and 2023 waves of the Youth Risk Behavior Survey for ages 14-18 (dashed lines) and Phase 3.2–Phase 4.2 of the Household Pulse Survey for ages 18-87 (solid lines). Survey weights are applied, 95% confidence intervals shaded. “Queer” refers to lesbian, gay, bisexual, and other non-straight identities.

Figure S8. Composition of kinship network by kin sexual identity and Focal age

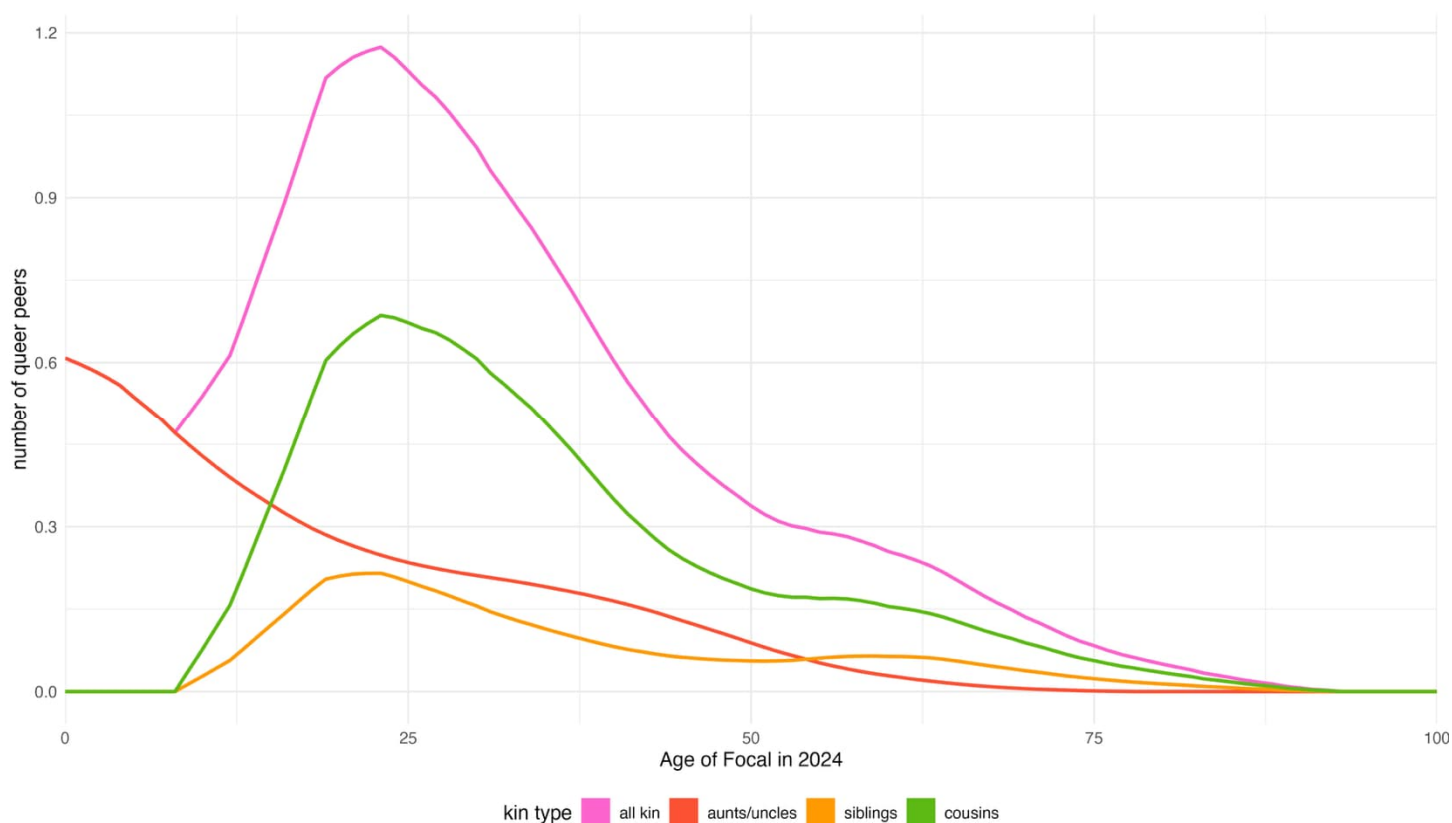
Notes: Estimates for an average person aged 15, 30, 45, 60, and 75 based on projected 2024 kinship structure and pooled 2021-2024 rates of queer identification. The corresponding birth cohort of Focal is written in parentheses beneath their age in 2024. Within each bar, the darker shade indicates the expected number of queer kin, the middle shade indicates the expected number of heterosexual (straight) kin, and the lighter shade captures “unknown” kin for whom we cannot estimate their identity due to being out of the age range for the survey. Queer includes lesbian/gay, bisexual, and other non-straight identities. “Nibling” = niece or nephew. *Source:* Estimates based on two-sex, time-variant kinship matrix model. Demographic data are drawn from the Human Mortality Database and Human Fertility Database. Survey data on queer identification are drawn from the 2021 and 2023 waves of Youth Risk Behavior Survey for ages 14-17 and Phase 3.2–Phase 4.2 of the Household Pulse Survey for ages 18-87. Survey weights are applied to ensure the findings are nationally representative.

Figure S9. Joint age-specific expected number of queer kin by age of Focal and age of kin, by kin type

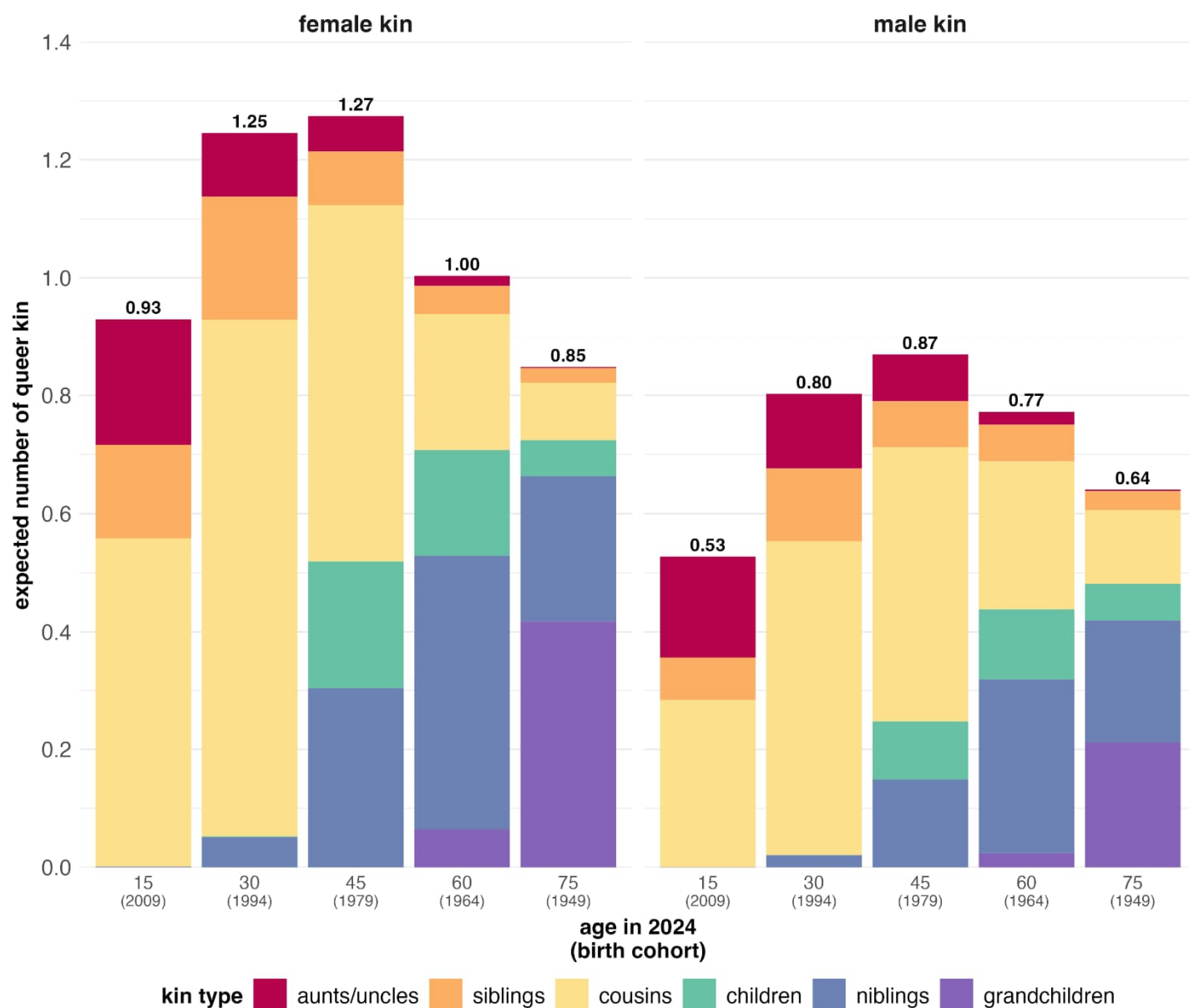


Notes: Lighter shades indicate relatively more queer kin, whereas darker shades indicate relatively fewer.

Figure S10. Distribution of queer peers, by age of Focal and kin type



Notes: Queer peers are individuals within a five-year age range of Focal, or 10 or more years older for aunts/uncles.

Figure S11. Expected number of queer female and male kin, by kin type and Focal age

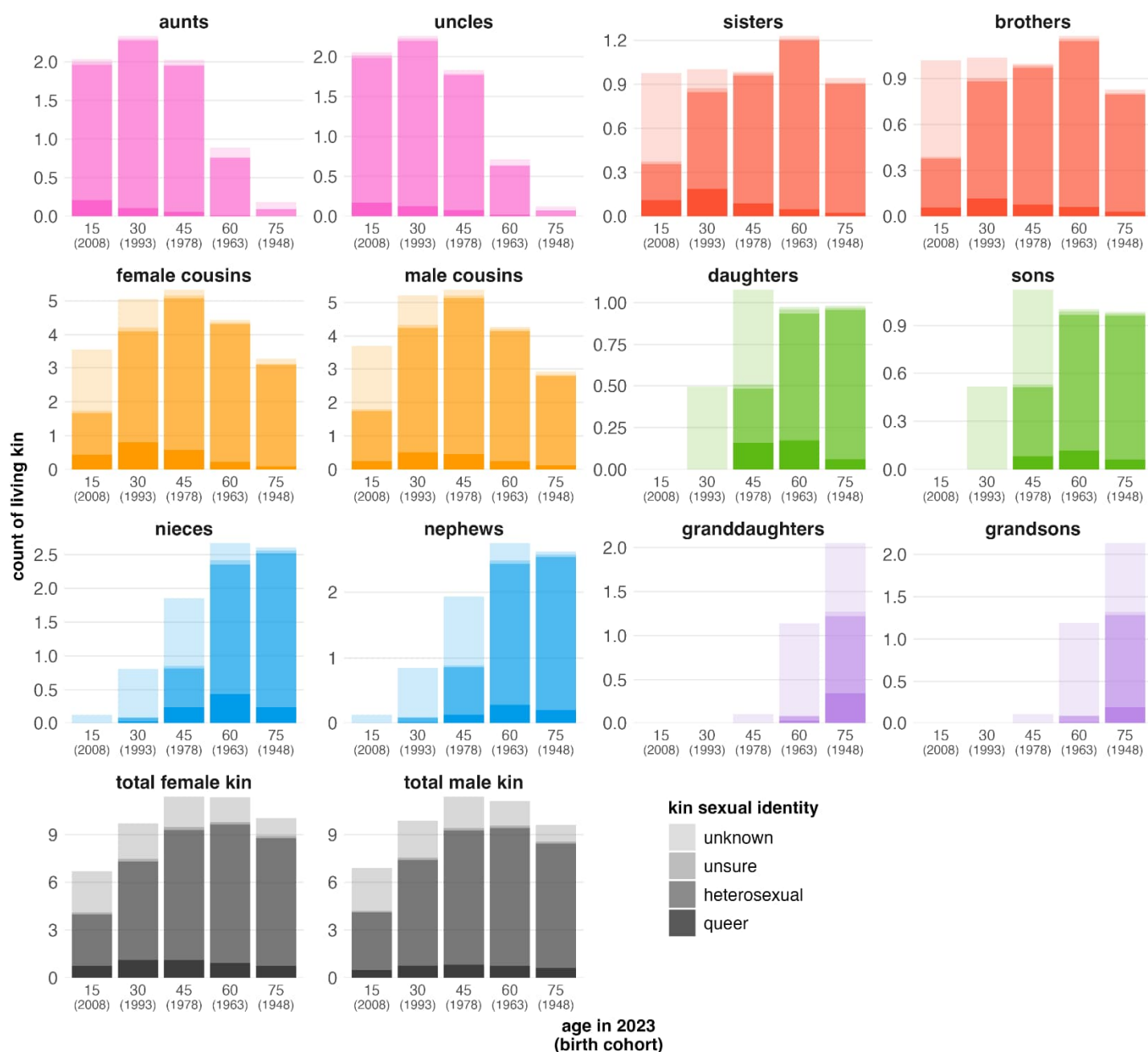
Source: Survey data on queer identification are drawn from the 2021 and 2023 waves of Youth Risk Behavior Survey for ages 14-17 and Phase 3.2–Phase 4.2 of the Household Pulse Survey for ages 18-87. Survey weights are applied to ensure the findings are nationally representative.

D. Additional Tables and Figures Corresponding to Findings

Table S3. Estimates corresponding to Figure 2

Age of Focal in 2024 (Cohort)	Unknown (LL, UL)	Unsure (LL, UL)	Straight (LL, UL)	Queer (LL, UL)
<i>Aunts/Uncles</i>				
15 (2009)	0.08 (0.07,0.09)	0.06 (0.05,0.08)	3.53 (3.49,3.57)	0.38 (0.35,0.41)
30 (1994)	0.07 (0.05,0.09)	0.05 (0.04,0.06)	4.18 (4.14,4.21)	0.23 (0.21,0.26)
45 (1979)	0.11 (0.1,0.13)	0.04 (0.02,0.05)	3.63 (3.6,3.65)	0.14 (0.12,0.16)
60 (1964)	0.21 (0.2,0.22)	0.01 (0.01,0.02)	1.4 (1.38,1.41)	0.04 (0.03,0.05)
75 (1949)	0.15 (0.15,0.15)	0 (0,0)	0.17 (0.17,0.17)	0 (0,0.01)
<i>Siblings</i>				
15 (2009)	1.22 (1.22,1.22)	0.03 (0.02,0.04)	0.57 (0.55,0.58)	0.17 (0.16,0.18)
30 (1994)	0.27 (0.26,0.27)	0.05 (0.04,0.06)	1.43 (1.41,1.46)	0.31 (0.28,0.33)
45 (1979)	0.02 (0.01,0.03)	0.03 (0.02,0.04)	1.74 (1.73,1.76)	0.17 (0.15,0.18)
60 (1964)	0.04 (0.03,0.05)	0.03 (0.02,0.03)	2.2 (2.18,2.21)	0.11 (0.1,0.12)
75 (1949)	0.06 (0.05,0.07)	0.02 (0.01,0.02)	1.7 (1.69,1.72)	0.06 (0.05,0.07)
<i>Cousins</i>				
15 (2009)	3.69 (3.68,3.7)	0.11 (0.09,0.14)	2.67 (2.62,2.73)	0.68 (0.62,0.73)
30 (1994)	1.69 (1.67,1.71)	0.21 (0.16,0.25)	6.95 (6.84,7.05)	1.29 (1.2,1.39)
45 (1979)	0.32 (0.29,0.35)	0.18 (0.14,0.22)	9.21 (9.12,9.31)	1.05 (0.96,1.13)
60 (1964)	0.16 (0.12,0.19)	0.1 (0.07,0.13)	8.03 (7.97,8.1)	0.48 (0.43,0.53)
75 (1949)	0.26 (0.23,0.29)	0.06 (0.04,0.08)	5.75 (5.7,5.8)	0.22 (0.19,0.25)
<i>Children</i>				
15 (2009)	0 (0,0)	0 (0,0)	0 (0,0)	0 (0,0)
30 (1994)	0.97 (0.97,0.97)	0 (0,0)	0 (0,0)	0 (0,0)
45 (1979)	1.15 (1.14,1.15)	0.04 (0.03,0.05)	0.73 (0.71,0.75)	0.24 (0.22,0.26)
60 (1964)	0.03 (0.03,0.04)	0.04 (0.03,0.05)	1.63 (1.61,1.65)	0.3 (0.28,0.32)
75 (1949)	0.02 (0.02,0.03)	0.02 (0.02,0.03)	1.75 (1.73,1.76)	0.12 (0.11,0.13)
<i>Niblings</i>				
15 (2009)	0.24 (0.24,0.24)	0 (0,0)	0 (0,0)	0 (0,0)
30 (1994)	1.46 (1.46,1.46)	0.01 (0.01,0.01)	0.12 (0.12,0.12)	0.04 (0.04,0.04)
45 (1979)	2.03 (2.03,2.03)	0.06 (0.05,0.07)	1.28 (1.25,1.31)	0.35 (0.33,0.38)
60 (1964)	0.5 (0.49,0.52)	0.11 (0.09,0.14)	4.01 (3.95,4.07)	0.72 (0.66,0.77)
75 (1949)	0.09 (0.08,0.11)	0.08 (0.06,0.1)	4.68 (4.63,4.72)	0.45 (0.41,0.49)
<i>Grandchildren</i>				
15 (2009)	0 (0,0)	0 (0,0)	0 (0,0)	0 (0,0)
30 (1994)	0 (0,0)	0 (0,0)	0 (0,0)	0 (0,0)
45 (1979)	0.19 (0.19,0.19)	0 (0,0)	0 (0,0)	0 (0,0)
60 (1964)	2.13 (2.13,2.13)	0.01 (0.01,0.01)	0.12 (0.11,0.12)	0.04 (0.04,0.04)
75 (1949)	1.58 (1.57,1.58)	0.09 (0.07,0.11)	1.89 (1.85,1.94)	0.52 (0.48,0.56)
<i>Total</i>				
15 (2009)	5.23 (5.21,5.25)	0.21 (0.16,0.25)	6.77 (6.66,6.88)	1.23 (1.13,1.32)
30 (1994)	4.46 (4.42,4.51)	0.31 (0.24,0.38)	12.68 (12.51,12.84)	1.87 (1.73,2.02)
45 (1979)	3.82 (3.75,3.88)	0.34 (0.26,0.43)	16.6 (16.41,16.79)	1.94 (1.78,2.11)
60 (1964)	3.07 (3,3.15)	0.3 (0.22,0.38)	17.38 (17.2,17.56)	1.68 (1.54,1.83)
75 (1949)	2.16 (2.09,2.23)	0.26 (0.19,0.34)	15.94 (15.77,16.11)	1.38 (1.25,1.51)

Notes: Authors' calculation based on data from the Household Pulse Survey. "LL" = lower limit to 95% confidence interval, "UL" = upper limit to 95% confidence interval.

Figure S12. Composition of kinship network by kin sexual identity, kin sex, and Focal age

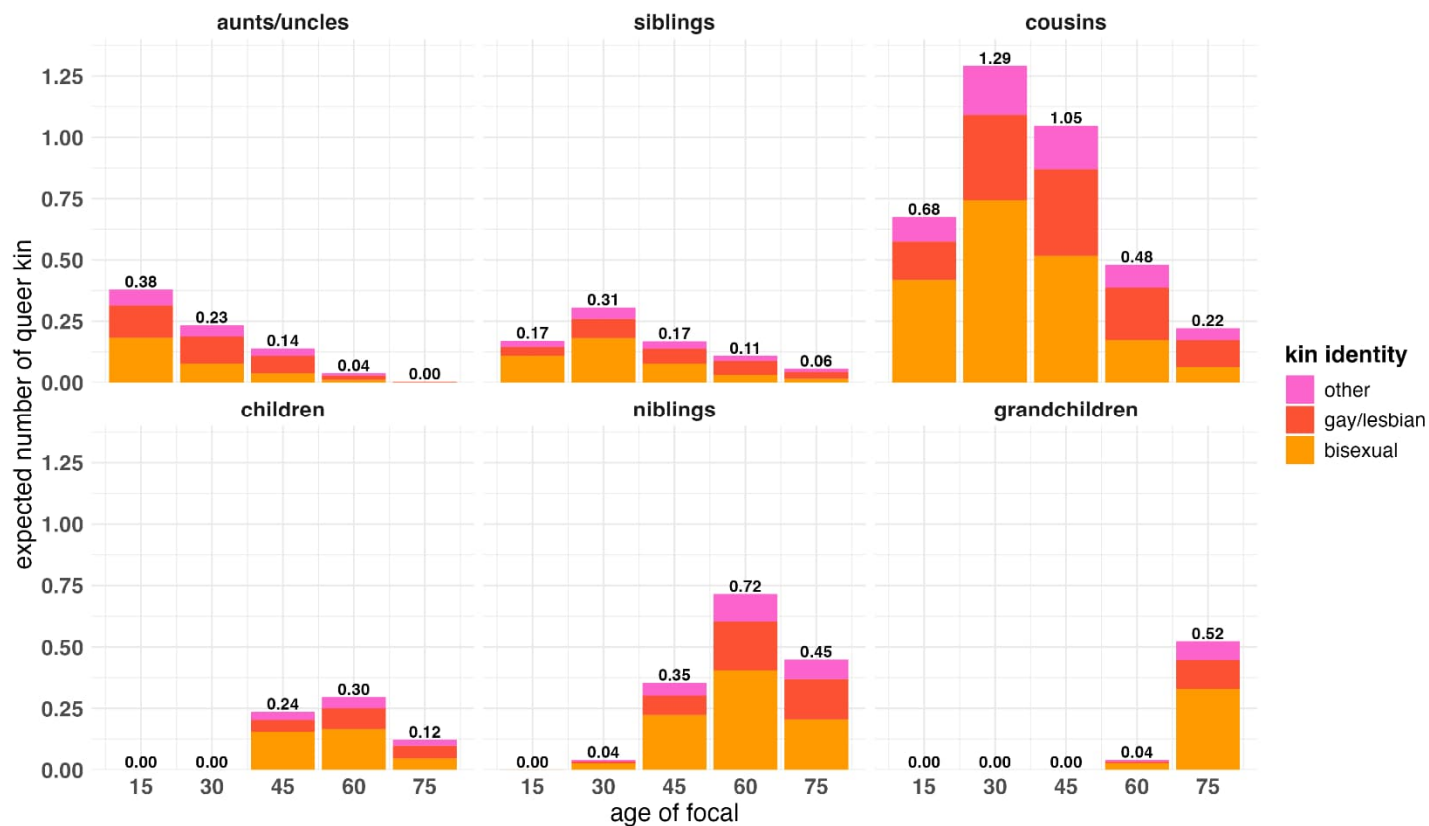
Notes: Estimates for an average person aged 15, 30, 45, 60, and 75 based on projected 2024 kinship structure and pooled 2021–2024 rates of queer identification. The corresponding birth cohort of Focal is written in parentheses beneath their age in 2024. Within each bar, the darker shade indicates the expected number of queer kin, the middle shade indicates the expected number of heterosexual (straight) kin, and the lighter shade captures “unknown” kin for whom we cannot estimate their identity due to being out of the age range for the survey. Queer includes lesbian/gay, bisexual, and other non-straight identities. “Nibbling” = niece or nephew. *Source:* Estimates based on two-sex, time-variant kinship matrix model. Demographic data are drawn from the Human Mortality Database and Human Fertility Database. Survey data on queer identification are drawn from Phase 3.2–Phase 4.2 of the Household Pulse Survey for ages 18–87. Survey weights are applied to ensure the findings are nationally representative.

Table S4. Estimates corresponding to Figure 5

Age in 2024 (Birth Cohort)		
Kin type	Female kin (LL, UL)	Male kin (LL, UL)
<i>15 (2009)</i>		
aunts/uncles	0.21 (0.2,0.22)	0.17 (0.15,0.19)
siblings	0.11 (0.1,0.12)	0.06 (0.05,0.07)
cousins	0.43 (0.4,0.46)	0.25 (0.22,0.27)
children	0 (0,0)	0 (0,0)
niblings	0 (0,0)	0 (0,0)
grandchildren	0 (0,0)	0 (0,0)
<i>30 (1994)</i>		
aunts/uncles	0.11 (0.1,0.12)	0.13 (0.11,0.14)
siblings	0.19 (0.18,0.2)	0.12 (0.11,0.13)
cousins	0.79 (0.74,0.83)	0.51 (0.46,0.55)
children	0 (0,0)	0 (0,0)
niblings	0.03 (0.02,0.03)	0.01 (0.01,0.02)
grandchildren	0 (0,0)	0 (0,0)
<i>45 (1979)</i>		
aunts/uncles	0.06 (0.05,0.07)	0.08 (0.07,0.09)
siblings	0.09 (0.08,0.1)	0.08 (0.07,0.09)
cousins	0.59 (0.55,0.62)	0.46 (0.41,0.5)
children	0.16 (0.15,0.17)	0.08 (0.07,0.09)
niblings	0.23 (0.21,0.24)	0.13 (0.11,0.14)
grandchildren	0 (0,0)	0 (0,0)
<i>60 (1964)</i>		
aunts/uncles	0.02 (0.01,0.02)	0.02 (0.02,0.03)
siblings	0.05 (0.04,0.05)	0.06 (0.05,0.07)
cousins	0.23 (0.21,0.25)	0.25 (0.22,0.28)
children	0.18 (0.17,0.19)	0.12 (0.11,0.13)
niblings	0.43 (0.41,0.46)	0.28 (0.26,0.31)
grandchildren	0.03 (0.03,0.03)	0.01 (0.01,0.02)
<i>75 (1949)</i>		
aunts/uncles	0 (0,0)	0 (0,0)
siblings	0.02 (0.02,0.03)	0.03 (0.03,0.04)
cousins	0.1 (0.08,0.11)	0.12 (0.11,0.14)
children	0.06 (0.06,0.07)	0.06 (0.06,0.07)
niblings	0.24 (0.23,0.26)	0.21 (0.19,0.23)
grandchildren	0.34 (0.32,0.36)	0.19 (0.17,0.21)

Notes: Authors' calculation based on data from the Household Pulse Survey. "LL" = lower limit to 95% confidence interval, "UL" = upper limit to 95% confidence interval.

Figure S13. Expected number of living female and male queer kin, by kin type, kin sexual identity, and Focal birth cohort



Notes: Estimates based on projected 2024 kinship structure and pooled 2021-2024 rates of queer identification.

Robustness check: Poisson approximation for share with 1+ queer kin

We also use an alternative measure of the share of one or more queer kin in which we instead assume that the number of queer kin is approximately Poisson distributed. Table S5 below shows that this alternative measure is substantively similar to the measure presented in the main text (see Table 1).

Table S5. Share of population with one or more queer kin, by kin sexuality, kin genealogical closeness, and Focal age using Poisson approximation

Age of Focal and Kin Sexuality	Any kin ^a (%, 95% CI)	Genealogical closeness	
		Nuclear kin ^b (%, 95% CI)	Extended kin ^c (%, 95% CI)
Focal aged 15 (Born in 2009)			
Queer ^d	70.6 (67.7, 73.3)	15.7 (14.5, 16.8)	65.2 (62.2, 67.9)
Bisexual	50.9 (47.2, 54.3)	10.4 (9.4, 11.4)	45.2 (41.7, 48.4)
Gay/Lesbian	27.8 (24, 31.3)	3.6 (2.9, 4.2)	25.1 (21.7, 28.3)
Other	17.3 (13.5, 20.9)	2.4 (1.8, 3)	15.2 (11.9, 18.4)
Focal aged 30 (Born in 1994)			
Queer	84.6 (82.2, 86.7)	26.3 (24.7, 28)	79.1 (76.4, 81.5)
Bisexual	64.2 (60.3, 67.8)	16.6 (15.1, 18)	57.1 (53.2, 60.7)
Gay/Lesbian	42.1 (37.3, 46.5)	7.5 (6.5, 8.6)	37.4 (33, 41.5)
Other	25.7 (20.4, 30.6)	4.5 (3.6, 5.5)	22.2 (17.5, 26.6)
Focal aged 45 (Born in 1979)			
Queer	85.7 (83.2, 87.8)	33.3 (31.1, 35.4)	78.6 (75.6, 81.1)
Bisexual	63.5 (59.3, 67.3)	20.6 (18.7, 22.5)	54.1 (49.9, 57.9)
Gay/Lesbian	45.8 (40.8, 50.5)	10.4 (8.9, 12)	39.5 (35, 43.7)
Other	27.6 (21.7, 33.1)	6.2 (4.7, 7.6)	22.8 (17.8, 27.6)
Focal aged 60 (Born in 1964)			
Queer	81.4 (78.5, 83.9)	33.3 (31.1, 35.4)	72.1 (68.8, 75)
Bisexual	55.7 (51.3, 59.7)	17.9 (16.2, 19.6)	46 (41.9, 49.8)
Gay/Lesbian	44 (39.2, 48.4)	13.1 (11.5, 14.7)	35.6 (31.3, 39.5)
Other	25 (19.3, 30.2)	6.5 (5, 8)	19.7 (15, 24.2)
Focal aged 75 (Born in 1949)			
Queer	74.8 (71.3, 77.8)	16.5 (14.8, 18.1)	69.8 (66.4, 72.9)
Bisexual	48.4 (43.9, 52.5)	6 (5, 7)	45.1 (41, 48.9)
Gay/Lesbian	37.8 (33.1, 42.1)	7.7 (6.6, 8.8)	32.5 (28.4, 36.4)
Other	21.5 (16.1, 26.6)	3.7 (2.6, 4.8)	18.5 (13.9, 22.9)

Notes: Estimates based on 2024 kinship structure projected by time-variant, two-sex kinship matrix model and pooled 2021-2024 rates of queer identification. Method relies on Poisson approximation. Kin are inclusive of the 18-87 age range. "CI" = 95% confidence interval. Categories are not mutually exclusive; estimates are irrespective of if, e.g., Focal has queer extended kin (for queer nuclear cell), or if Focal has bisexual nuclear kin (for gay/lesbian extended kin cell). ^a Nuclear V extended kin. ^b Siblings V children, exclusive of parents and partners. ^c aunts/uncles V cousins V nieces/nephews V grandchildren. ^d Bisexual V gay/lesbian V other non-straight.

E. Subnational queer schedules

Although focusing on queer kinship at the subnational level is outside the scope of this study, we present four figures (see Figures S14-S17) to demonstrate queer schedules at various subnational levels: college attainment, region of residence, year, and race-ethnicity. Here, standard errors (estimated using replicate weights) are often larger and estimates far less smooth due to smaller sample size at each age-sex-level combination. These figures generally show similarities between subgroups, which we interpret to mean that any differences in the experience of queer kinship at the subnational level would be due to differences in kinship size and structure, rather than patterns of queer identification.

Figure S14. Queer schedule, by college attainment

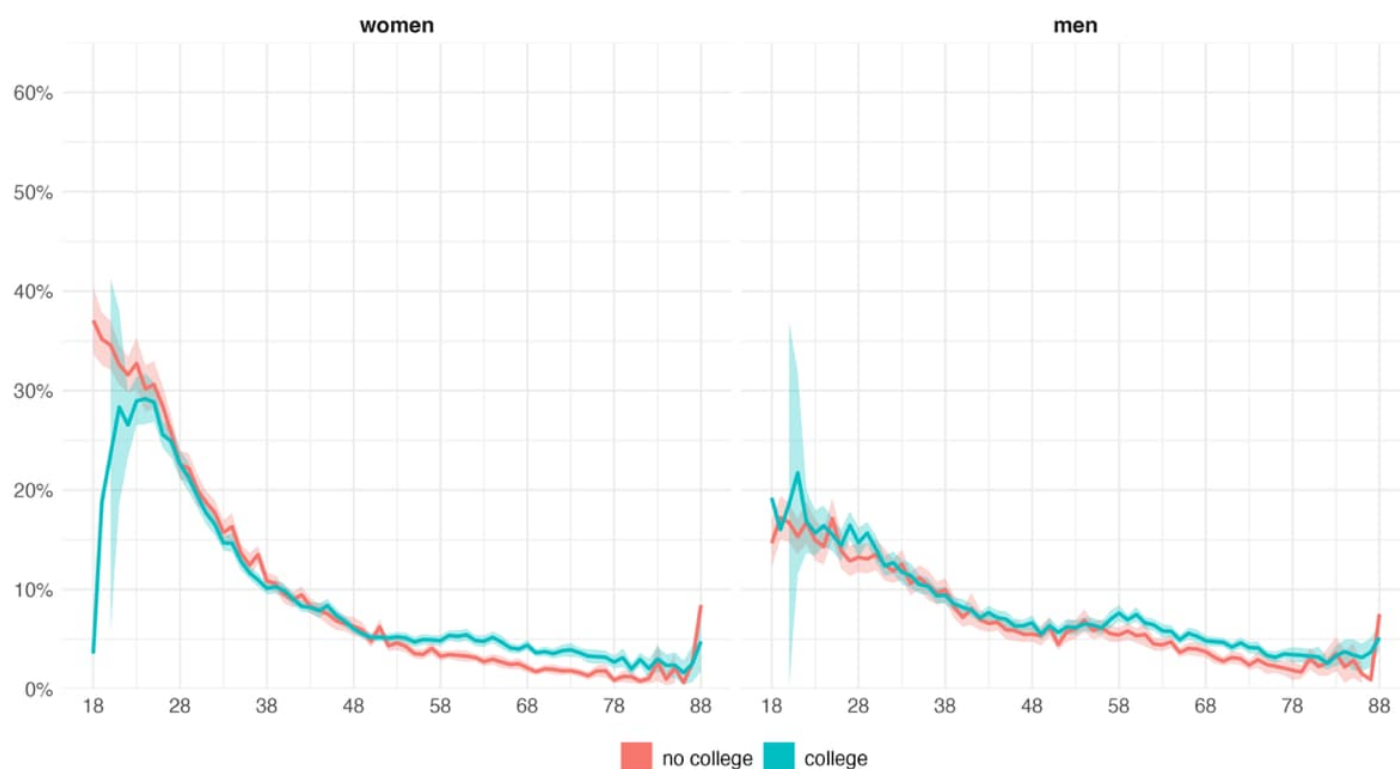


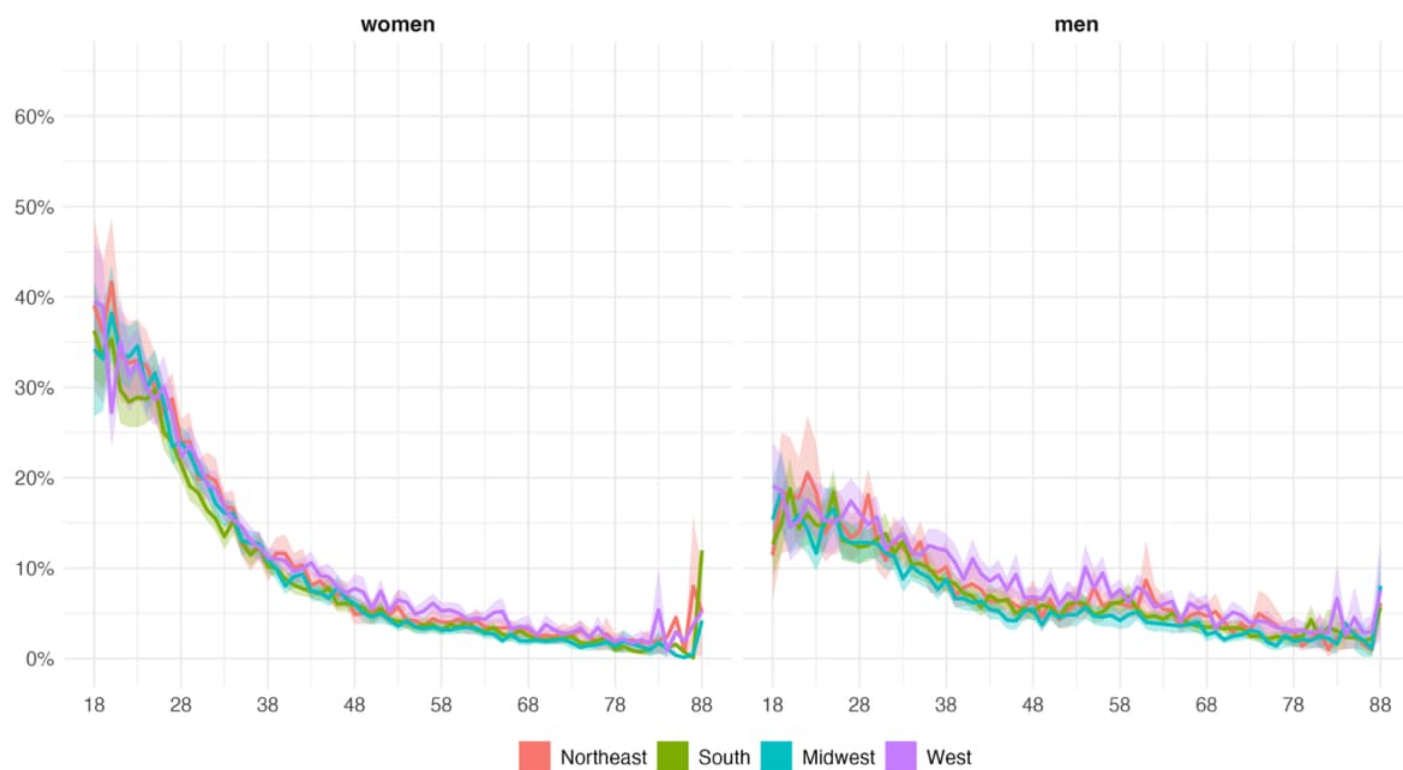
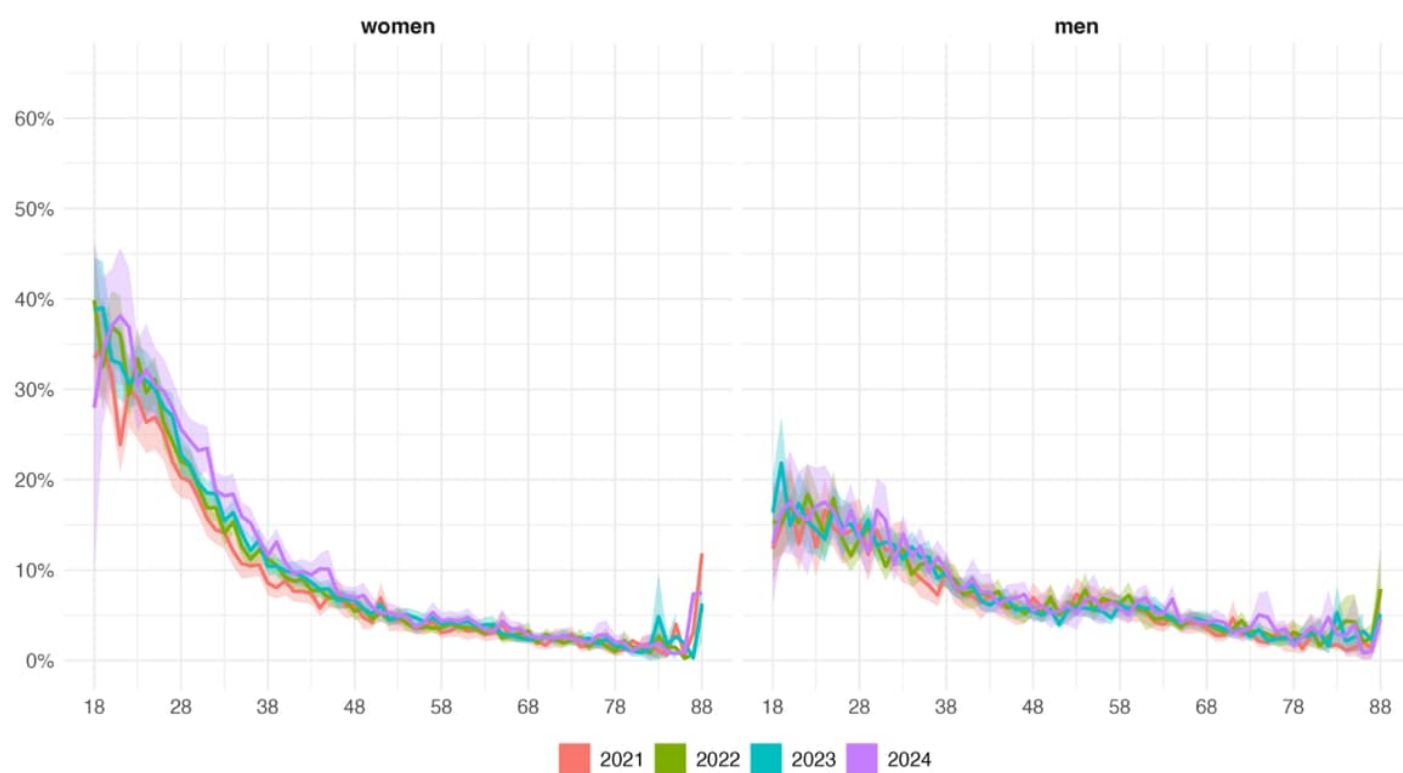
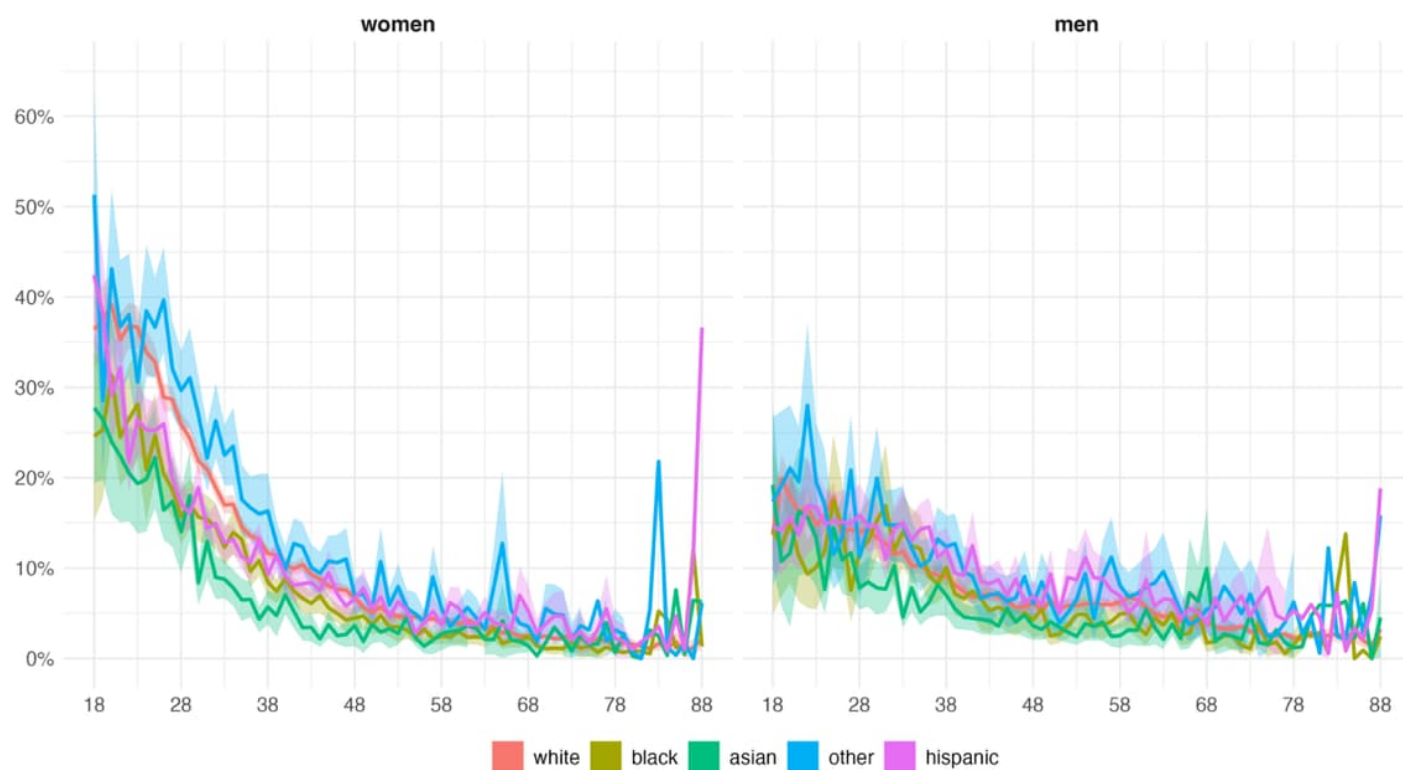
Figure S15. Queer schedule, by region**Figure S16.** Queer schedule, by year

Figure S17. Queer schedule, by race-ethnicity

F. Replication

Readers wishing to replicate our findings can do so via the following link: [removed for peer review].

G. References

- Abrams, L., Jiang, S., Alburez-Gutierrez, D., Acosta, E., & Zagheni, E. (2025). *Adverse Mortality Trends have Increased the Number Bereaved in the United States*. SocArXiv. https://doi.org/10.31235/osf.io/z38cs_v1
- Alburez-Gutierrez, D., Williams, I., & Caswell, H. (2023). Projections of human kinship for all countries. *Proceedings of the National Academy of Sciences*, 120(52), e2315722120. <https://doi.org/10.1073/pnas.2315722120>
- Brady, L. A., Julian, C. A., & Manning, W. D. (2025). Variation Between LGBT Estimates and State Policy Context. *Population Research and Policy Review*, 44(1), 11. <https://doi.org/10.1007/s11113-025-09938-2>
- Deng, B., & Watson, T. (2023). *LGBTQ+ Data Availability: What We Can Learn From Four Major Surveys*. Brookings Institute. https://www.brookings.edu/wp-content/uploads/2023/04/20230417_CCF_DengWatson_MAF-1.pdf
- Hacker, J. D., & Roberts, E. (2020). Fertility decline in the United States, 1850-1930: New evidence from complete-count datasets: *Annales de Démographie Historique*, n° 138(2), 143–177. <https://doi.org/10.3917/adh.138.0143>
- Julian, C. A., Manning, W. D., & Westrick-Payne, K. K. (2024). Responses to Sexual and Gender Identity Measures in Population-Level Data by Birth Cohort: A Research Note. *Demography*, 61(1), 15–30. <https://doi.org/10.1215/00703370-11164985>
- Lee, S., Fredriksen-Goldsen, K. I., McClain, C., Kim, H.-J., & Suzer-Gurtekin, Z. T. (2018). Are Sexual Minorities Less Likely to Participate in Surveys? An Examination of Proxy Nonresponse Measures and Associated Biases with Sexual Orientation in a Population-Based Health Survey. *Field Methods*, 30(3), 208–224. <https://doi.org/10.1177/1525822X18777736>
- Macapagal, K., Coventry, R., Arbeit, M. R., Fisher, C. B., & Mustanski, B. (2017). “I Won’t Out Myself Just to Do a Survey”: Sexual and Gender Minority Adolescents’ Perspectives on the Risks and Benefits of Sex Research. *Archives of Sexual Behavior*, 46(5), 1393–1409. <https://doi.org/10.1007/s10508-016-0784-5>
- Schoumaker, B. (2019). Male Fertility Around the World and Over Time: How Different is it from Female Fertility? *Population and Development Review*, 45(3), 459–487. <https://doi.org/10.1111/padr.12273>